Research on the Relationship between Entrepreneur Confidence Index and Producer Price Index Based on Quantile Granger Causality

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Abstract

This paper investigates the relationship between Entrepreneur Confidence Index and Producer Price Index (PPI) based on the mean and quantile Granger causality tests, using quarterly statistics data from 2005 to 2017. The results indicate that there is a unidirectional causality between entrepreneur confidence index and PPI index. At different quantile levels, entrepreneur confidence index of the previous period has different effects on the current PPI index. At lower quantiles, the causality from entrepreneur confidence index to PPI index is significant, while entrepreneur confidence index has little impact on PPI at higher quantiles. The estimates of $\hat{\beta}_i$ at low quantiles are positive and significant. Therefore, improving entrepreneur investment confidence will have a certain positive impact on PPI.

Keywords

Entrepreneur Confidence Index, Producer Price Index, Quantile Regression, Granger Causality Test

1. Introduction

China's economic development has entered a new normal at present, the regional economy and industrial structure have been continuously improved, and the economy operation maintains a steady and positive trend. Since the second half of 2016, entrepreneur confidence index constantly rallied and rose 4.5% year-on-year in the first half of 2016, the production and development market is active, and the efficiency of enterprises is improved. Entrepreneurs think that
current macro-control policy will have a positive effect on future economy. China’s producer price index rose 6.4% year-on-year over the period January 2017 to July; China has achieved a good momentum while at the same time securing progress in its economic development. This may be related to market participant’s expectations of economic conditions in the future and good prospects of the economy. Entrepreneurs are the leaders of enterprise production and operation, whose performance and behavior often play a vital role, influencing economic development significantly. In the General Theory of Employment, Interest, and Money [1], Keynes pointed out that when people were confident about the future economic development, they would increase investment to expand the scale of production and operation; when people lacked confidence in the future economic development, they would reduce production capacity to minimize the losses, thereby inhibiting investment and consumption. Therefore, market confidence will affect the economic development by influencing the decision-making of economic actors. With the above in mind, this paper takes producer price index (PPI) as the research object, and adopts the quarterly entrepreneur confidence index as the measurement index of entrepreneur confidence, which is published by Oriental Fortune Net. We analyze and empirically investigate the relationship between entrepreneur confidence and PPI to provide corresponding policy suggestions for improving the confidence of entrepreneurs and forecasting the trend of economic development.

There are many types of models that have been used for researching entrepreneur confidence and PPI. Xu et al. (2010) used the Grey forecasting model (GM(1, 1)) to analyze China’s entrepreneur confidence index [2]. Zhang (2012) introduced the artificial neural network model for forecasting the PPI index by using China’s factory price index data in 1990-2008 [3]. Liu et al. (2015) investigated the predictive effect of entrepreneur confidence index by applying time relation analysis and Granger causality [4]. Cheng et al. (2015) adopted VAR model to forecast PPI index [5]. Dong et al. (2016) analyzed and predicted the PPI from January 2014 to July 2014 based on ARIMA model [6]. Luo et al. (2018) introduced a multiple conduction model for forecasting PPI index [7]. Many scholars have also done a lot of research on the relationship among entrepreneur confidence, PPI and other economic variables. Pan et al. (2010) established VAR model and BVAR model to explore the relationship between entrepreneur confidence and inflation [8]. Qianlong (2016) investigated the relationship between entrepreneur confidence index and PMI based on quantile regression [9]. Gui et al. (2016) introduce the time frequency analysis method EEMD-JADE into economic field to study the structure analysis and transmission mechanism of the relationship between PMI and PPI [10]. Song et al. (2016) employed ARDL-ECM model to study the driving factors affecting market confidence, their findings suggested that stock price index, GDP and CPI could boost the entrepreneur confidence in the long-run [11]. Shikai et al. (2016) examined the relationship between CPI and PPI using VEC model and concluded that CPI had a long-term
equilibrium relation with PPI [12]. Nie et al. (2017) established the VAR model to measure the relationship among China’s agricultural product price index, PPI and economic growth [13]. Luo et al. (2018) used VAR-VEC model to analyze the relationship between consumer confidence index (CCI) and PPI [14].

Previous studies are more common to test non-causality in conditional mean based on a linear model, it reveals the average influence in different economic variables, while cannot depict the influence at other places, we are therefore motivated to characterize and test causal relations in other conditional distribution characteristics. In this paper, we analyze the effect of entrepreneur confidence index on PPI index using both the mean and quantile Granger causality tests, and the causal relations between them were discussed. The rest of the study is organized as follows. We introduce the notion of Granger causality in mean and quantiles in Section 2. The empirical results of different causal models are presented in Section 3. Section 4 concludes the study.

2. Empirical Methodology

2.1. Mean Granger Causality Test

Granger (1969) proposed a causal relation based on prediction [15], which mainly examined whether the explanatory variable X was in favor of predicting Y, if X and its lag values were increased, the prediction ability of Y conditional mean could be significantly enhanced. We define that X is the Granger cause of Y, call it \( X \Rightarrow Y \). Otherwise, we say that X is not the Granger cause of Y and call it \( X \not\Rightarrow Y \). The Granger causality is used as the null hypothesis by defining non-causality between variables, namely,

\[
F_{\eta} \left[ \eta \mid (Y, X)_{t-1} \right] = F_{\eta} \left[ \eta \mid Y_{t-1} \right], \quad \forall \eta \in \mathbb{IR}
\]

where \( F_{\eta} \left[ \cdot \mid F \right] \) is the conditional distribution of \( Y_t \), and \((Y, X)_{t-1}\) is the information set generated by \( Y_t \) and \( X_t \) up to time \( t-1 \). The variable \( X \) is said to Granger causes \( Y \) when Equation (1) fails to hold. As estimating and testing conditional distribution are practically cumbersome, it is more common to test a necessary condition of (1), namely,

\[
E \left[ Y_t \mid (Y, X)_{t-1} \right] = E \left[ Y_t \mid Y_{t-1} \right]
\]

where \( E \left[ Y_t \mid F \right] \) is the mean of \( F_{\eta} \left[ \cdot \mid F \right] \). We say that \( X \) does not Granger cause \( Y \) in mean if (2) holds; otherwise, \( X \) Granger causes \( Y \) in mean.

In order to test whether the entrepreneur confidence index (QYJ) is the mean Granger cause of PPI index, we consider null hypothesis \( H_0 : \text{QYJ} \not\Rightarrow \text{PPI} \) or \( H_0 : \beta_0 = \cdots = \beta_q = 0 \), the two unconstrained and constrained mean regression models are needed to be built:

\[
E (\text{PPI}_t) = a_0 + \sum_{i=1}^{p} \alpha_i \text{PPI}_{t-i}
\]

\[
E (\text{PPI}_t) = a_0 + \sum_{i=1}^{p} \alpha_i \text{PPI}_{t-i} + \sum_{j=1}^{q} \beta_j \text{QYJ}_{t-j}
\]
where $PPI_t$ is the producer price index at time $t$, $i$ is the lag order of producer price index, $p$ is the maximum lag order of producer price index, $QYJ_t$ is the entrepreneur confidence index at time $t$, $j$ is the lag order of entrepreneur confidence index, $q$ is the maximum lag order of entrepreneur confidence index, $QYJ_{t-j}$ is the entrepreneur confidence index at time $(t-j)$, $a_0$ is constant term. Equation (3) is a constrained mean regression model, which indicates that $QYJ$ has no significant influence on $PPI$; Equation (4) is a unconstrained mean regression model, the independent variable of the model depends on the lagged values of $PPI_t$ and $QYJ_t$, which indicates $QYJ$ has a significant effect on $PPI$.

We construct F test statistics according to the residual sum of squares of both constrained and unconstrained mean regression models, namely,

$$F = \frac{(ESS_1 - ESS_2)/p}{ESS_2/[T-(q+p+1)]} - F(p,T-(q+p+1))$$

(5)

where $ESS_1$ is the residual sum of squares of constrained model, $ESS_2$ is the residual sum of squares of unconstrained model. If the statistic $F > F_a$ ($a$ is a confidence), then we reject the null hypothesis and it can be considered that $QYJ$ is the mean Granger cause of $PPI$, otherwise the null hypothesis is accepted.

When using Granger causality, time series are required to be stationary or non-stationary with cointegration relations, otherwise the results will be wrong. In addition, as causality test is sensitive to lag order, several lag orders are usually tested in actual experiment. If the test results are consistent, the results are considered to be more reliable.

### 2.2. Quantile Granger Causality Test

The mean Granger causality test is limited to test the causal relations between the mean value of explanatory variable on the conditional distribution of response variable, and it is unable to describe its predictive ability of response variable at other places. The quantile Granger causality test can make up for this deficiency. By combining the quantile regression with the mean Granger causality test, it can find the specific location where the causal relationship between variables is established, so as to catch the causal relationship that cannot be found by the mean Granger causality test. Therefore, other conditional distribution characteristics are considered to test the causal relationship. The quantile Granger causality not only can test the causality on the mean value of conditional distribution, but also can test the causality of variables on different conditional quantiles, it can fully describe the overall picture of the conditional distribution of response variables. And the quantile regression does not require a strong hypothesis for the error term, so the estimator of quantile regression coefficient is more robust for the non-normal distribution.

Given that a distribution is completely determined by its quantile, Granger non-causality in distribution can also be expressed in terms of conditional quantiles [16]. Letting $Q_{\tau} (\tau \cdot |\cdot)$ denotes the $\tau$-th quantile of $F_{\gamma} [\cdot |\cdot]$, (1) is equivalent to
We say that $X$ does not Granger cause $Y$ in all quantiles if (6) holds. We may define Granger non-causality in a quantile range $(b,c) \subset (0,1)$ as

$$Q_n[\tau | (Y, X)_{i,b,c}] = Q_n[\tau | Y_{i,b,c}], \ \forall \tau \in (b,c) \quad (7)$$

To test whether $X$ is the quantile Granger cause of $Y$, the linear quantile regression model is established as follow:

$$y_\tau = a_\tau (\tau) + \alpha (\tau) y_{\tau-1,p} + \beta (\tau) x_{\tau-1,q} + \ell_\tau (\tau) = z_{\tau-1}^{\prime} \theta (\tau) + \ell_\tau (\tau) \quad (8)$$

where $\theta (\tau) = \begin{bmatrix} a_\tau (\tau), \alpha (\tau), \beta (\tau) \end{bmatrix}$ is the k-dimensional parameter vector with $k = 1 + p + q$, and $\ell_\tau (\tau)$ is the corresponding error.

Given a linear model for conditional quantiles, testing (6) amounts to testing $H_0 : \beta (\tau) = 0, \ \forall \tau \in (0,1) \quad (9)$

For a given $\tau$, it can be derived the Wald statistic of $\beta (\tau) = 0$ through Equation (8), that is:

$$W_{T(\tau)} = T \hat{\beta}_{T(\tau)}^\prime (\psi \hat{M}_T \psi) \cdot \hat{\beta}_{T(\tau)} \tau (1-\tau)$$

(10)

where $\hat{\beta}_{T(\tau)}$ is the estimate of $\beta (\tau)$, $\psi$ is a selection matrix, and the formula $\psi \theta (\tau) = \beta (\tau)$ is hold, $\hat{M}_T = \frac{1}{T} \sum_{t=1}^{T} Z_{\tau-1} Z_{\tau-1}^\prime$, $\hat{j}$ is probability density function.

To test (9), Koenker and Machado (1999) suggested using a Sup-Wald test, the supremum of $W_{T(\tau)}$.

In what follows, let $B_{q}$ denote a vector of $q$ independent Brownian bridges, $\Rightarrow$ denote weak convergence (of associated probability measures) and $\| \|$ denote the Euclidean norm. Clearly, $B_{q} (\tau)$ equals $\left[ \tau (1-\tau) \right]^{1/2} N \left( 0, I_{q} \right)$ in distribution, and the formula as follow:

$$\sqrt{T} \left[ \hat{\beta}_{T(\tau)} (\tau) - \beta (\tau) \right] \Rightarrow \Omega^{1/2} B_{q} (\tau) \quad (11)$$

Under suitable conditions, (11) holds uniformly on a closed interval $\Lambda \subset (0,1)$; see Koenker and Machado (1999) for details. We then have, under the null hypothesis (9),

$$W_{T(\tau)} \Rightarrow \left\| \frac{B_{q} (\tau)}{\sqrt{\tau (1-\tau)}} \right\|^2, \ \tau \in \Lambda$$

(12)

where the weak limit is the sum of squares of $q$ independent Bessel processes. This immediately leads to the following result:
ANDREWS [17] and KOENKER [18] calculated the critical values of sup-Wald statistics according to the simulation method. Table 1 presents the simulated critical values of sup-Wald test (with $q = 1, 2$) on $[0.05, 0.95]$.

3. Empirical Study

3.1. Data Description

Entrepreneur confidence index (QYJ) reflects entrepreneur’s feeling and confidence in macroeconomic environment, it is an indicator to forecast the changing trend of economic development. The value is between 0 and 200 with 100 as the critical value, a reading above 100 indicates a rise in confidence, the economy is in a state of prosperity. While one below 100 suggests that confidence has decreased, the economy is in recession. The PPI is the barometer of national economic conditions, which reflects the prosperity and depression of the economy. In order to intuitively understand the changing relationship between entrepreneur confidence index and PPI index over years, this paper collects quarterly statistical data from the first quarter 2005 to the fourth quarter 2017, the changing trend of the entrepreneur confidence index and PPI index is plotted from the data, as shown in Figure 1.

As can be seen from Figure 1, the volatility of the entrepreneur confidence index during the sample period is greater than that of the PPI index, but the changing trend of entrepreneur confidence index is basically consistent with PPI index. That is, the entrepreneur confidence index rises and the PPI rises, conversely,

<table>
<thead>
<tr>
<th>$q$</th>
<th>$a = 0.1$</th>
<th>$a = 0.05$</th>
<th>$a = 0.01$</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>8.19</td>
<td>9.84</td>
<td>13.01</td>
</tr>
<tr>
<td>2</td>
<td>11.20</td>
<td>12.93</td>
<td>16.44</td>
</tr>
</tbody>
</table>

Table 1. Critical values of sup-Wald test ($q = 1, 2$).

Figure 1. The changing trend of the entrepreneur confidence index and PPI index.
the entrepreneur confidence index falls and so does PPI. The changes between the two series can be mainly divided into three parts: one is from 2005 to the end of 2007, China’s economy is overheating in this period, and both entrepreneur confidence index and PPI index are running at a high level, the entrepreneur confidence index is much higher than 100, indicating that enterprises are positive and optimistic about the current economic situation and future expectation. Macroeconomic investment rises and the economy has entered a rapid growing stage; Two is from 2008 to 2010, influenced by international financial crisis and in the economic downturn, entrepreneurs lack confidence in investment. China’s economic development has received a greater negative impact, especially in the fourth quarter of 2008, the entrepreneur confidence index is always below 100; Three is the post-crisis from 2010 to 2017, the government has issued a series of economic policies to strengthen domestic and foreign investment to stimulate economic growth. The confidence of enterprise investors has gradually recovered. However, due to the lack of technological innovation capacity, structural industry surplus and other problems, the growth momentum of China’s economic development is not strong, and the entrepreneur confidence index has dropped significantly.

Figure 2 and Figure 3 depict the normality test results of the PPI index series. In Figure 2, the histogram shows obvious bimodal characteristics, which indicates that the PPI index series is not normal distribution, and there is a large deviation between its kernel density curve (solid line) and normal density curve (dotted line). The upper tail of the Q-Q diagram deviates significantly from the
In Figure 3, we therefore reject the assumption that the PPI index series follows normal distribution. This result shows that the precondition for mean regression model based on the classical assumption is no longer true. For this purpose, we need to establish a linear quantile regression model to explore the relationship between entrepreneur confidence index and PPI index.

3.2. Data Source and Processing

To study the causality between the entrepreneur confidence index and PPI index, the quarterly data from 2005 to 2017 are selected as variables. Among them, the entrepreneur confidence index (denoted as QYJ) is obtained from Oriental Fortune Net (http://www.eastmoney.com/), the logarithmic form is expressed by LNQYJ, and the first order difference is DLNQYJ. The PPI index data is obtained from National Bureau of Statistics (http://data.stats.gov.cn/), the seasonal average of PPI is expressed as LNPPPI in logarithmic form and DLNPPI in first difference. Besides, in order to eliminate the effect of heteroscedasticity on regression results and reduce volatility, the data used in the empirical part were all taken with the natural logarithm, and the transformation of natural logarithm will not change the characteristics of the original data.

3.3. Unit Root Test

The Augmented Dickey-Fuller (ADF) unit root test is employed to test stationar-
ity of the series. Checking for stationarity of data series is an important prerequisite in most empirical time series analysis, as these methods require stationarity of the variables. Results of unit root test are reported in Table 2. The results show that ADF statistics of LNQYJ and LNPPI are more than the critical values of 1%, 5% and 10%, and the p-values are also more than 0.05, so we cannot reject the null hypothesis of unit roots for both variables in level form. However, the null hypothesis is rejected when ADF unit root test is applied to the first differences of each variable. The first differences of the QYJ and PPI are stationary indicating that these variables are integrated of order one, I (1).

3.4. Cointegration Test

The purpose of cointegration test is to prevent spurious regression. There are two theories to test the cointegration relationship of time series. One is Engle-Granger (E-G) two step method, the other is Johansen test based on VAR model. This paper uses E-G two step method to test the cointegration relationship between LNQYJ and LNPPI series. The E-G two step method is to conduct unit root test of regression residual series based on OLS model. Firstly, we apply OLS to fit equation

$$y_t = \alpha + \beta x_t + \mu_t$$

and the regression equation is used to calculate the non-equilibrium error \(\hat{\mu}_t\); then the stationarity of residual series \(\hat{\mu}_t\) is tested. If \(\hat{\mu}_t\) is a stationary series, it is considered that \(x_t\) and \(y_t\) are cointegration variables, and there is a long-run equilibrium relationship between them. As Engle and Granger (1987) pointed out, only non-stationary variables with the same order of integration could be tested for cointegration. Since QYJ and PPI series are integrated with the same order I (1), cointegration test can be conducted by E-G two step method.

By using OLS estimation, the cointegration regression equation is

$$LNPPI = 85.37 + 0.13LNQYJ + \mu_{t}$$

$$= (21.50062) (3.404283)$$

(13)

The t statistics of each parameter are shown within parentheses in Equation (13), the adjusted R-squared is 0.865107, the F-statistics is 75.35606, the p-values of the constant term C and the explanatory variable LNQYJ are 0.0000 and 0.0014 respectively, both of which are less than 0.05. It can be seen from the regression results that the fitting effect of the cointegration regression equation is very good.

Table 2. Results of ADF unit root test.

<table>
<thead>
<tr>
<th>Series</th>
<th>ADF test value</th>
<th>Critical value</th>
<th>P value</th>
<th>Conclusion</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td></td>
<td>1%</td>
<td>5%</td>
<td>10%</td>
</tr>
<tr>
<td>LNQYJ</td>
<td>-2.785002</td>
<td>-3.56831</td>
<td>-2.9212</td>
<td>-2.898551</td>
</tr>
<tr>
<td>LNPPI</td>
<td>0.019370</td>
<td>-2.61301</td>
<td>-1.9477</td>
<td>-1.612573</td>
</tr>
<tr>
<td>DLNQYJ</td>
<td>-5.906184</td>
<td>-2.61203</td>
<td>-1.9475</td>
<td>-1.61265</td>
</tr>
<tr>
<td>DLNPPi</td>
<td>-4.933082</td>
<td>-2.61301</td>
<td>-1.9477</td>
<td>-1.612573</td>
</tr>
</tbody>
</table>
And then the unit root test is performed on the estimated residual series $\hat{\mu}_t$. The critical value applied by ADF test during the cointegration analysis is different from the traditional ADF test, but refers to the table of cointegration critical values provided by Engle-Granger. The equation for calculating critical value is $C(\alpha) = \phi_c + \phi_1 T^{-1} + \phi_2 T^{-2}$, where $T$ represents the sample size and $\alpha$ is the significance level. According to Equation (13), the critical value $C(\alpha)$ is $-3.3410$ at the given $\alpha = 0.05$. When conducting ADF test of residual series, the test statistic is $-4.531120$, which is less than the critical value at $\alpha = 0.05$. The result shows that we can reject the null hypothesis. Therefore, the residual series is stationary. So we have reasons to believe that, there is a cointegration relationship between LNQYJ and LNPPPI, and there is a long-run equilibrium relationship between the entrepreneur confidence index and PPI index. In the long-run, every 1% rise in entrepreneur confidence index is associated with increases of 0.13% in PPI index.

Based on the above ADF test and E-G cointegration test, we find that the entrepreneur confidence index and PPI index are both non-stationary time series, and there is a long-run equilibrium relationship between the two series, thus it is suitable for the two economic variables to analyze their relationship by using Granger causality test.

### 3.5. Mean Granger Causality Test

The Granger causality test is used to analyze the causal relationship between the variables’ conditional mean, which is to test whether the explanatory variable $X$ is favorable for predicting the behavior of the prerequisite is that $X$ and $Y$ are stationary series or non-stationary series with cointegration relations. Granger pointed out that if two integral series with the same order had cointegration relationship, there must be causality to support this long-run equilibrium, and the influence might be one-way or two-way. To explore the relationship between the entrepreneur confidence index and PPI index, the mean Granger causality test is performed on LNQYJ and LNPPPI series. Since the optimal lag order of VAR is one, this paper chooses the mean Granger causality test of one lag order. In order to make the test results more reliable, the two lag order and three lag order are also tested. **Table 3** presents the results of mean Granger causality test.

<table>
<thead>
<tr>
<th>Null hypothesis</th>
<th>Observation</th>
<th>Lag order</th>
<th>F-statistic</th>
<th>P value</th>
<th>Conclusion</th>
</tr>
</thead>
<tbody>
<tr>
<td>LNPPPI $\rightarrow$ LNQYJ</td>
<td>51</td>
<td>1</td>
<td>3.61883</td>
<td>0.0631</td>
<td>accept</td>
</tr>
<tr>
<td>LNQYJ $\rightarrow$ LNPPPI</td>
<td>51</td>
<td>1</td>
<td>14.7861</td>
<td>0.0004</td>
<td>reject</td>
</tr>
<tr>
<td>LNPPPI $\rightarrow$ LNQYJ</td>
<td>50</td>
<td>2</td>
<td>0.94621</td>
<td>0.3958</td>
<td>accept</td>
</tr>
<tr>
<td>LNQYJ $\rightarrow$ LNPPPI</td>
<td>50</td>
<td>2</td>
<td>5.1057</td>
<td>0.01</td>
<td>reject</td>
</tr>
<tr>
<td>LNPPPI $\rightarrow$ LNQYJ</td>
<td>49</td>
<td>3</td>
<td>1.33439</td>
<td>0.276</td>
<td>accept</td>
</tr>
<tr>
<td>LNQYJ $\rightarrow$ LNPPPI</td>
<td>49</td>
<td>3</td>
<td>3.38919</td>
<td>0.0266</td>
<td>reject</td>
</tr>
</tbody>
</table>
From Table 3, at 5% level, the results show that the PPI index is not the Granger cause of the entrepreneur confidence index, while reject the entrepreneur confidence index is not the Granger cause of PPI index. It indicates that the change in entrepreneur confidence index can cause the change in PPI index to some extent.

3.6. Quantile Granger Causality Test

For the quantile Granger causality test, the causality between variable LNPPI, and LNQYJ, can be tested in two ways: one is to test a single coefficient, that is, the significance of the lag parameter of the independent variable is tested respectively. We considered the null hypothesis $H_0: \beta_j = 0 (j = 1, 2, \cdots, q)$. LNQYJ is considered as the Granger causal of LNPPI, if the hypothesis is not true at a certain $\tau$. The other is the joint test of all parameters, the null hypothesis is $H_0: \beta_1 = \cdots = \beta_q = 0$. LNQYJ is considered as the Granger causal of LNPPI, if the hypothesis is not true at all $\tau$.

3.6.1. Single Coefficient Test in Quantile

For each PPI-QYJ relation, we consider the following model:

$$y_i = \alpha(\tau) + \sum_{\tau_1} a_{\tau_1}(\tau) y_{i,\tau_1} + \sum_{\tau_2} \beta_{\tau_2}(\tau) x_{i,\tau_2} + \mu_i(\tau)$$ (14)

To determine whether PPI index Granger causes entrepreneur confidence index, $y$ is LNPPI and $x$ is LNQYJ; for reversed causal relations, $y$ is LNQYJ and $x$ is LNPPI. This model specification allows us to investigate whether lagged $x$ delivers information (about $y$) that is not contained in lagged $y$. To illustrate, we estimate model (14) with $p = q = 1$ according to the optimal lag order of VAR. For each $y$, 19 quantile regressions (with $\tau = 0.05, 0.1, \cdots, 0.9, 0.95$) least-squares regression (OLS) are computed. The results of Granger causality test in quantile are reported in Table 4.

The coefficient estimates $\hat{\beta}_j$, the test statistics $t$ and the p-values are provided for this test. Table 4 shows the testing of whether PPI index Granger causes entrepreneur confidence index. Since the p-values are greater than 0.05 across entire conditional distribution (when $0.05 \leq \tau \leq 0.95$), it is evident that there is no causality from PPI index to entrepreneur confidence index, the result is the same as OLS test and the above mean Granger causality test. We conclude that PPI index changes do not lead entrepreneur confidence index changes in $0.05 \leq \tau \leq 0.95$.

When we test whether entrepreneur confidence index Granger causes PPI index, the p-values are significant and are all less than 0.05 in $0.05 \leq \tau \leq 0.45$, hence there are strong reasons to accept that entrepreneur confidence index Granger causes PPI index at lower quantiles. However, when $0.5 \leq \tau \leq 0.95$, it is not significant that entrepreneur confidence index Granger causes PPI index, which indicates entrepreneur confidence index has little impact on PPI index when the conditional distribution of PPI is located at higher quantile. The result of the OLS test also show that there is causality from entrepreneur confidence index to
Table 4. Results of quantile Granger causality test.

<table>
<thead>
<tr>
<th>τ</th>
<th>QYJ Granger causes PPI</th>
<th></th>
<th>PPI Granger causes QYJ</th>
<th></th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>( \hat{\beta} )</td>
<td>t</td>
<td>P</td>
<td>( \hat{\beta} )</td>
</tr>
<tr>
<td>0.05</td>
<td>0.2249</td>
<td>4.1352</td>
<td>0.0001</td>
<td>−1.9622</td>
</tr>
<tr>
<td>0.1</td>
<td>0.1881</td>
<td>4.6161</td>
<td>0.0000</td>
<td>−0.7935</td>
</tr>
<tr>
<td>0.15</td>
<td>0.1767</td>
<td>4.4311</td>
<td>0.0001</td>
<td>−0.2885</td>
</tr>
<tr>
<td>0.2</td>
<td>0.1828</td>
<td>4.5138</td>
<td>0.0000</td>
<td>−0.3307</td>
</tr>
<tr>
<td>0.25</td>
<td>0.1907</td>
<td>4.4539</td>
<td>0.0001</td>
<td>0.3029</td>
</tr>
<tr>
<td>0.3</td>
<td>0.2057</td>
<td>4.3532</td>
<td>0.0001</td>
<td>0.2348</td>
</tr>
<tr>
<td>0.35</td>
<td>0.1699</td>
<td>3.0362</td>
<td>0.0039</td>
<td>0.1934</td>
</tr>
<tr>
<td>0.4</td>
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</table>

PPI index, while this test is only conducted at a mean level and does thus not provide an overall picture of the existing causality from entrepreneur confidence index to PPI index. Quantile Granger causality test can provide the influence of entrepreneur confidence at different quantiles. Through this test, we not only look at the causality beyond the mean estimates, but we also account for the structural breaks.

It can be seen from Table 4 that the regression estimates of \( \hat{\beta} \) vary with quantiles. The values of \( \hat{\beta} \) at lower quantiles are greater than that at higher quantiles. They are significantly positive for lower quantiles but insignificant at higher quantiles (\( \tau \) in \([0.5, 0.95]\)). This implies that, the effect of entrepreneur confidence on PPI has an obvious difference when PPI is in different periods. For the PPI index at low quantiles, the estimates of \( \hat{\beta} \) on the previous entrepreneur confidence index are positive and significant, we therefore conclude that improving entrepreneur investment confidence will have a certain positive impact on PPI.
3.6.2. Joint Test in Quantiles

To be sure, this paper applies the sup-Wald test to check joint significance of $\hat{\beta}_i$ on $[0.05, 0.95]$. We first test the Granger causality from entrepreneur confidence index to PPI index, the sup-Wald statistic is 10.3 and reject the null hypothesis of estimate $\hat{\beta}_i = 0$ at 5% level, this indicates that entrepreneur confidence Granger cause PPI. We then analyze the causality from PPI index to entrepreneur confidence index, the sup-Wald statistic is 5.6 and cannot reject the null hypothesis of estimate $\hat{\beta}_i = 0$ at 5% level, thus there is no evidence to believe that the PPI index is the Granger causality for entrepreneur confidence index. The results of the joint test are the same as the single coefficient test, except that the single coefficient test know the causality in which areas of the quantile.

4. Conclusion

In this study, cointegration, and methodology of Granger causality test are employed to empirically investigate causal link between entrepreneur confidence index and PPI index in China. We make use of quarterly data from the first quarter 2005 to the fourth quarter 2017. The cointegration test results indicate that there is a long-run equilibrium relationship between the entrepreneur confidence index and PPI index. And in the long-run, every 1\% rise in entrepreneur confidence index is associated with increases of 0.13\% in PPI index. Both the mean and quantile causality test indicate a one-way causality between the entrepreneur confidence index and PPI index. The direction of causality is from entrepreneur confidence index to PPI index. The impact of entrepreneur confidence on PPI varies with quantiles. At lower quantiles, the causality from entrepreneur confidence index to PPI index is significant, and plays a positive role in promoting PPI. While entrepreneur confidence index has little impact on PPI at higher quantiles. The estimates of $\hat{\beta}_i$ at low quantiles are positive and significant. The results indicate that improving entrepreneur investment confidence will have a certain positive impact on PPI. When analyzing whether PPI index is the causality for entrepreneur confidence index, the p-values are not significant either in the mean test or in the quantile test. It fails to pass the significant test at the 5% level. The result indicates that PPI index is not the Granger causality for entrepreneur confidence index.

Conflicts of Interest

The authors declare no conflicts of interest regarding the publication of this paper.

References


