Option Pricing Applications of Quadratic Volatility Models

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ABSTRACT

Recently there has been a surge of interest in higher order moment properties of time varying volatility models. Various GARCH-type models have been developed and successfully applied in empirical finance. Moment properties are important because the existence of moments permit verification of how well theoretical models match stylized facts such as fat tails in most financial data. In this paper, we consider various types of random coefficient autoregressive (RCA) models with quadratic generalized autoregressive conditional heteroscedasticity (GARCH) errors and study the moments, mean, variance and kurtosis. We also consider the Black-Scholes model with RCA GARCH volatility and show that these moments can be used to evaluate the call price for European options.

Keywords: Garch Processes; RCA Models; Garch Models; Time Varying Volatility; Kurtosis

1. Introduction

It is well-known that many financial time series such as stock returns exhibit leptokurtosis and time-varying volatility [1]. The generalized autoregressive conditional heteroscedasticity (GARCH) and the random coefficient autoregressive (RCA) models have been extensively used to capture the time-varying behaviour of the volatility. Studies using GARCH models commonly assume that the time series is conditionally normally distributed; however, the kurtosis implied by the normal GARCH tends to be lower than the sample kurtosis observed in many time series Bollerslev [1]. Thavaneswaran et al. [2] use an ARMA representation to derive the kurtosis of various classes of GARCH models such as power GARCH, non-Gaussian GARCH, non-stationary and random coefficient GARCH. Recently, Thavaneswaran et al. [3], Appadoo et al. [4] have extended the results to stationary RCA processes with GARCH errors and Paseka et al. [5] further extended the results to RCA processes with stochastic volatility (SV) errors.

Leptokurtosis is commonly observed in financial time series, as well as in currency and commodity markets. The opening and closure of the markets, time-of-the-day and day-of-the-week effects, weekends and vacation periods cause changes in the trading volume that translates into regular changes in price variability. Financial, currency, and commodity data also respond to new information entering into the market, which usually have large kurtosis. Recently, there has been growing interest in using volatility models [3,4]. Most of the studies use GARCH models with dummy variables in the volatility equation, and a few of them have been extended to a more flexible form such as the RCA GARCH. However, even though much research has been performed on volatility models applied to market data such as stock returns, more general specifications accounting for RCA with GARCH errors have been little explored. First we derive the kurtosis of a simple time series model with behaviour in the mean. Then we introduce various classes of RCA GARCH models and study the moments and discuss applications in option pricing. We extend the results for RCA GARCH volatility models to RCA quadratic GARCH models. The RCA GARCH model is appropriate for time series where significant autocorrelation exists. Option pricing with RCA model with quadratic GARCH errors is also discussed in some detail. The moments derived for the RCA GARCH volatility models provide more accurate estimates of market data behaviour and help investors, decision makers, and other market participants develop improved trading strategies. The rest of the paper is organized as follows. In rest of Section 1, we present results on standard GARCH models. These results are interesting for their own sake. In Section 2, we derive the higher order moments of some RCA models with GARCH errors, and in Section 3 we discuss some option pricing applications with RCA models with GARCH errors.

GARCH Models

Consider the general class of GARCH \((P,Q)\) model for
Consider the class of random coefficient autoregressive (RCA) models defined by allowing random additive perturbations of the autoregressive (AR) coefficients of ordinary AR models. That is, we assume that the process \( y_t \) is given by,

\[
y_t - \sum_{i=1}^{p} (\phi_i + b_i(t))y_{t-i} = e_i
\]

where \( Z_t \) is a sequence of independent, normally distributed random variables with zero mean, unit variance. Let \( u_i = e_i^2 - h_i \) be the martingale difference and let \( \sigma_i^2 \) be the variance of \( u_i \). (1.2) and (1.2) could be written as:

\[
\dot{\varepsilon}_i^2 - u_i = \omega + \sum_{j=1}^{p} \alpha_j \dot{\varepsilon}_{i-j}^2 + \sum_{j=1}^{q} \beta_j h_{i-j}
\]

and \( R = \max(P, Q) \). We shall make the following stationarity assumptions for \( \dot{\varepsilon}_i^2 \) which has an ARMA\((R, Q)\) representation. All the zeroes of the polynomial \( \Phi(B) \) lie outside of the unit circle.

\[
\sum_{i=0}^{\infty} \Psi_i^2 < \infty \text{ where the } \Psi_i \text{s are obtained from the relation } \psi(B) = \Phi(B) \text{ with }
\]

\[
\Psi(B) = 1 + \sum_{i=1}^{\infty} \Psi(B^i)
\]

The assumption sensures that the \( \dot{u}_i \)'s are uncorrelated with zero mean and finite variance and that the \( \dot{\varepsilon}_i^2 \) process is weakly stationary. In this case, the autocorrelation function of \( \dot{\varepsilon}_i^2 \) will be exactly the same as that for stationary ARMA\((R, Q)\) model. For any random variable \( \varepsilon \) with finite fourth moments, the kurtosis defined by

\[
\frac{E((\varepsilon - \mu)^4)}{Var(\varepsilon)}
\]

is normal then the process \( \{e_i\} \) defined by Equations (1.3) and (1.4) is called a normal GARCH \((p, q)\) process. The kurtosis of the GARCH process is denoted by \( K(\varepsilon) \) when it exists.

2. Random Coefficient Volatility Models

Consider the class of random coefficient autoregressive (RCA) models defined by allowing random additive perturbations of the autoregressive (AR) coefficients of ordinary AR models. That is, we assume that the process \( y_t \) is given by,

\[
y_t - \sum_{i=1}^{p} (\phi_i + b_i(t))y_{t-i} = e_i
\]

where the parameters \( \theta_i, i = 2, \cdots, p \), are assumed to be known, \( e_i \) and \( b_i(t) \) are zero mean square integrable independent processes and the variances are denoted by \( \sigma_e^2 \) and \( \sigma_b^2 \). \( h_i(t) \) \((i = 1, 2, \cdots)\)'s are independent of \( e_i \) and \( y_{t-i} \) and may be thought of as incorporating structural changes. In order to motivate nonlinear forecasts for nonlinear models, we consider a class of estimating functions of the form

\[
g_n = \sum_{i=1}^{n} b_i \cdot h_i \quad [6],
\]

where \( h_i = y_i - E[y_i | F_{i-1}^n] = y_i - \sum_{j=1}^{p} \phi_j y_{i-j} \) and \( b_{i,j} \) is a function of \( y_{1}, y_{2}, \cdots, y_{i} \) and possibly the known parameters \( \phi_1, \cdots, \phi_p \) (i.e. we assume that the fitted model is available). If we restrict ourselves to a class of estimating functions of the above form then we can forecast the future value of \( y_{n+1} \) based on the observed values \( y_1, y_2, \cdots, y_n \) as \( \hat{y}_n(1) = E[y_{n+1} | y_1, y_2, \cdots, y_n] \). That is, whether we have an AR\((p)\) model or RCA\((p)\) model we will get the same linear predictor of \( y_{n+1} \). However, for the RCA model under consideration, we have

\[
E[y_i | F_{i-1}^n] = \sum_{i=1}^{n} \phi_i y_{i-j} \quad \text{and var}[y_i | F_{i-1}^n] = \sigma^2 + \sum_{j=1}^{k} \gamma^2 + \sigma_b^2.
\]

Thus, the conditional variance is a nonlinear function and hence the RCA model may be viewed as a non-linear time series model. Nicholls and Quinn [7] studied linear as well as some nonlinear (proposed) forecast by fitting a nonlinear (RCA) model for the classical lynx cycle data. Using heuristic reasoning they proposed a nonlinear forecast and \( \hat{y}_{n+1} = \text{sgn}(\phi_{n+1} y_n) \left( \phi_{n+1} y_n + \sigma_{n+1}^2 \right)^{1/2} \) they showed empirically that the forecast \( y_{n+1} \) is a better predictor (having smaller forecast errors when compared with the actual observations) than the linear forecast for the lynx data. It is of interest to note that by defining

\[
h_i = y_i^2 - E[y_i^2 \mid F_{i-1}^n],
\]

the optimal forecast for \( y_{n+1} \) can be obtained as

\[
y_{n+1}^* = \left[ E[y_i^2 \mid F_{i-1}^n] \right]^{1/2} = \text{sgn}(\phi_{n+1} y_n) \left[ \phi_{n+1} y_n^2 + \sigma^2 + \sigma_b^2 \right]^{1/2}.
\]

That is, the estimating function method can be used to obtain a nonlinear forecast for a nonlinear models by considering a class of elementary martingale estimating functions generated by nonlinear functions of the observations. Using a similar argument we could also obtain forecasts for various class of GARCH models, see Thavaneswaran and Heyde [6] for details. The main message is that RCA models could be used to improve the forecasting performance of stochastic volatility models.

**Lemma 2.1.** When \( Z_t \) is a standard normal random variable such that \( Z_t \sim N(0,1) \) then,
Now if \( t \in \mathbb{Z} \), then where \( Z \) denotes the set of integers and
\[ 1) \quad \left( b_t, \varepsilon_t \right) \sim \left( \begin{array}{cc} 0 & 0 \\ 0 & \sigma^2_b \end{array} \right), \]
\[ 2) \quad \phi^2 + \sigma^2_b < 1. \]

The sequences \( \{b_t\} \) and \( \{\varepsilon_t\} \) respectively, are the errors in the model.

**Theorem 2.1.** Let \( \{y_t\} \) be a modified RCA (1) time series with an absolute value random coefficient satisfying conditions (1) and (2). The modified RCA (1) model is given by
\[ y_t = (\phi + b_t)y_{t-1} + \varepsilon_t, \quad t \in \mathbb{Z}, \]
where \( \mathbb{Z} \) denotes the set of integers and
\[ 1) \quad \left( b_t, \varepsilon_t \right) \sim \left( \begin{array}{cc} 0 & 0 \\ 0 & \sigma^2 \end{array} \right), \]
\[ 2) \quad \phi^2 + \sigma^2 < 1. \]

The sequences \( \{b_t\} \) and \( \{\varepsilon_t\} \) respectively, are the errors in the model.

**Proof:**

The autocovariance and the autocorrelation functions are given by
\[ \gamma_k = \phi + \left( \frac{2b \sigma^2_b}{\sqrt{\pi}} \right) \Gamma \left( \frac{\delta + 1}{2} \right) \gamma_{k-1} \quad \text{and} \quad p_k = \phi + \left( \frac{2\sigma^2_b}{\sqrt{\pi}} \right) \Gamma \left( \frac{\delta + 1}{2} \right). \]

Thus, we have
\[ E(y_t) = 0, \quad E(y_t^2) = \phi^2 E(y_{t-1}^2) + 2\phi E(y_{t-1}^2) E(|b_t|) + E(y_{t-1}^2) E(|b_t|^2) + E(\varepsilon_t^2). \]
and we have

$$E(y_t^4) = \frac{3\sigma^4}{1 - 8\sqrt{2}\sigma^3_{\phi} \phi - 4\sigma^3_{\phi} - 4\phi^3\sigma_b \sqrt{\frac{2}{\pi}} - 6\phi^2\sigma^2_b} \left[ \phi^2 + 2\phi \frac{\sqrt{2}}{\pi} \sigma_b + \sigma^2_b \right]$$

$$+ \frac{6\sigma^4}{1 - 8\sqrt{2}\sigma^3_{\phi} \phi - 4\sigma^3_{\phi} - 4\phi^3\sigma_b \sqrt{\frac{2}{\pi}} - 6\phi^2\sigma^2_b} \left[ 1 - \left( \phi^2 + 2\phi \frac{\sqrt{2}}{\pi} \sigma_b + \sigma^2_b \right) \right]$$

$$= \frac{3\sigma^4}{1 - 8\sqrt{2}\sigma^3_{\phi} \phi - 4\sigma^3_{\phi} - 4\phi^3\sigma_b \sqrt{\frac{2}{\pi}} - 6\phi^2\sigma^2_b} \left[ 1 + \phi^2 + 2\phi \frac{\sqrt{2}}{\pi} \sigma_b + \sigma^2_b \right]$$

$$K^{(y)} = \frac{E[y_t^4]}{E[y_t^2]^2} = \left[ \frac{3\left( 1 - \left( \phi^2 + 2\phi \frac{\sqrt{2}}{\pi} \sigma_b + \sigma^2_b \right) \right)^2}{1 - 8\sqrt{2}\sigma^3_{\phi} \phi - 4\sigma^3_{\phi} - 4\phi^3\sigma_b \sqrt{\frac{2}{\pi}} - 6\phi^2\sigma^2_b} \right]$$

When $\sigma^2 = 0$, the kurtosis of the process $y_t$ converge to $K^{(y)} = 3$. Thus, the autocorrelation function is given by

$$E(y_t) = E((\phi + |b|)) y_{t-1} + \epsilon^2_t = 0 \quad (2.6)$$

$K^{(y)} = \frac{E[y_t^4]}{E[y_t^2]^2} = \left[ \frac{3\left( 1 - \left( \phi^2 + 2\phi \frac{\sqrt{2}}{\pi} \sigma_b + \sigma^2_b \right) \right)^2}{1 - 8\sqrt{2}\sigma^3_{\phi} \phi - 4\sigma^3_{\phi} - 4\phi^3\sigma_b \sqrt{\frac{2}{\pi}} - 6\phi^2\sigma^2_b} \right] \quad (2.4)$$

where we use the fact that $\rho_0 = 1$.

**Theorem 2.2.** Suppose $y_t$ is an Random Coefficient Moving average process model of the form

$$y_t = (\phi + |b_t|) y_{t-1} + \epsilon^2_t \quad (2.5)$$

where $b_t$ is an uncorrelated Gaussian process with zero mean and with variance $\sigma^2_b$. $\epsilon_t$ is an uncorrelated Gaussian process with zero mean and with variance $\sigma^2$. Then, we have the following

**Proof:**

$$E(y_t^4) = \frac{12\phi \sigma_b \sqrt{\frac{2}{\pi}} 3\sigma^2_b \sigma^2 + 6\sigma^2_b \sigma^2 + 6\phi^3 \sigma^2 \sigma^2 \sigma^2}{1 - \phi^4 - 4\phi \sqrt{\frac{2}{\pi}} \sigma^2_b - 6\phi^2 \sigma^2_b - 4\phi \sqrt{\frac{2}{\pi}} \frac{2}{\pi} - 3\sigma^2_b}$$

$$+ \frac{36\phi \sigma_b \sqrt{\frac{2}{\pi}} 6 \sigma^2 \sigma^2 + 18\sigma^2 \sigma^2 + 18\phi \sigma^2}{1 - \phi^4 - 8\phi \sqrt{\frac{2}{\pi}} \sigma^2_b - 6\phi^2 \sigma^2_b - 4\phi \sqrt{\frac{2}{\pi}} \frac{2}{\pi} - 3\sigma^2_b}$$

$$E(y_{t-1}^4) + 105\sigma^2 \sigma^2$$

$$\quad \left[ 1 - \phi^4 - 4\phi \sqrt{\frac{2}{\pi}} \sigma^2_b - 6\phi^2 \sigma^2_b - 4\phi \sqrt{\frac{2}{\pi}} \frac{2}{\pi} - 3\sigma^2_b \right]$$

$$= \frac{315\sigma^2}{1 - \phi^4 - 8\phi \sqrt{\frac{2}{\pi}} \sigma^2_b - 6\phi^2 \sigma^2_b - 4\phi \sqrt{\frac{2}{\pi}} \frac{2}{\pi} - 3\sigma^2_b}$$

$$E(y_{t-1}^2) + 315\sigma^2$$

$$\quad \left[ 1 - \phi^4 - 8\phi \sqrt{\frac{2}{\pi}} \sigma^2_b - 6\phi^2 \sigma^2_b - 4\phi \sqrt{\frac{2}{\pi}} \frac{2}{\pi} - 3\sigma^2_b \right]$$
Thus we have

\[
E(y_{t-1}^4) = \frac{36\phi\sigma_b^2}{2\pi} \left( \frac{2}{\pi} \sigma_e^6 + 18\phi^2 \sigma_e^6 + 18\phi^3 \sigma_e^6 \right) \left[ 1 - \left( \phi^2 + 2\phi\sigma_b \sqrt{\frac{2}{\pi} + \sigma_b^2} \right) \right] \left[ \frac{3\sigma_b^6}{1 - (\phi^2 + 2\phi\sigma_b \sqrt{\frac{2}{\pi} + \sigma_b^2})} \right] \\
+ \left( 1 - \phi^4 - 8\phi \left( \frac{2}{\pi} \sigma_b^3 - 6\phi^2 \sigma_b^3 - 2\phi^3 \sigma_b \sqrt{\frac{2}{\pi} - 3\sigma_b^4} \right) \right) 315\sigma_e^{12}
\]

The kurtosis of the process is given by

\[
K^{(s)} = \frac{36\phi\sigma_b^2}{2\pi} \left( \frac{2}{\pi} \sigma_e^6 + 18\phi^2 \sigma_e^6 + 18\phi^3 \sigma_e^6 \right) \left[ 1 - \left( \phi^2 + 2\phi\sigma_b \sqrt{\frac{2}{\pi} + \sigma_b^2} \right) \right] \left[ \frac{3\sigma_b^6}{1 - (\phi^2 + 2\phi\sigma_b \sqrt{\frac{2}{\pi} + \sigma_b^2})} \right] \\
+ \left( 1 - \phi^4 - 8\phi \left( \frac{2}{\pi} \sigma_b^3 - 6\phi^2 \sigma_b^3 - 2\phi^3 \sigma_b \sqrt{\frac{2}{\pi} - 3\sigma_b^4} \right) \right) 35 \left( \phi^2 + 2\phi\sigma_b \sqrt{\frac{2}{\pi} + \sigma_b^2} \right)^2 \\
+ \left( 1 - \phi^4 - 8\phi \left( \frac{2}{\pi} \sigma_b^3 - 6\phi^2 \sigma_b^3 - 2\phi^3 \sigma_b \sqrt{\frac{2}{\pi} - 3\sigma_b^4} \right) \right) 35 \left( \phi^2 + 2\phi\sigma_b \sqrt{\frac{2}{\pi} + \sigma_b^2} \right)^2
\]

(2.8)

When \( \sigma_b^2 = 0 \), the kurtosis of the process \( y_t \) converge to \( K^{(s)} = \frac{35 - 29\phi^2}{(1 + \phi^2)} \) and when \( \sigma_b^2 = 0 \), and \( \phi = 0 \) the kurtosis of the process \( y_t \) turns out to be 35.

**Theorem 2.3.** Suppose \( y_t \) is an Random Coefficient Moving average process model of the form

\[
y_t = \epsilon_t + (\phi |b_t|) \epsilon_{t-1}
\]

where \( b_t \) is an uncorrelated noise process with zero mean and with variance \( \sigma_b^2 \). \( \epsilon_t \) is an uncorrelated noise process with zero mean and with variance \( \sigma_e^2 \). Then, we have the following relationships

\[
E\left(y_t^2\right) = \sigma_e^4 \left( 1 + \phi^2 + 2\phi\sigma_b \sqrt{\frac{2}{\pi} + \sigma_b^2} \right)
\]

\[
E\left(y_t^4\right) = \sigma_e^4 \left[ 3\phi^4 + 3\sigma_b^4 + 1 \right] + 12\left( \frac{2}{\pi} \right) (1 + 2\sigma_b^2 + \phi^2) \phi \sigma_b + 6 \left( \sigma_b^2 + \phi^2 + 3\phi^2 \sigma_b^2 \right)
\]

and the autocorrelation functions are given by

\[
\rho_k = \begin{cases} 
1 & k = 0 \\
\frac{\phi + \sqrt{\frac{2}{\pi} \sigma_b^2}}{1 + \phi^2 + 2\phi\sigma_b \sqrt{\frac{2}{\pi} + \sigma_b^2}} & k = 1 \\
0 & k = 2, 3, \ldots
\end{cases}
\]

(2.10)
Proof:

\[ E(y_t^2) = E\left[\varepsilon_t^2 + 2\varepsilon_t\varepsilon_{t-1}\phi + 2\varepsilon_t\varepsilon_{t-1}\theta^2 + 2\varepsilon_t^2\phi|b_t| + \varepsilon_t^2\phi^2 + 2\varepsilon_t^2\theta^2 + 2\varepsilon_t^2|b_t|^2\right], \]

\[ = \sigma_e^2 + \phi^2\sigma_e^2 + 2\phi\sigma_e^2\sigma_h^2\frac{2\pi}{\pi} + \sigma_e^2\sigma_h^2 = \sigma_e^2\left(1 + \phi^2 + 2\phi\sigma_h^2\frac{2\pi}{\pi} + \sigma_h^2\right). \]

\[ E(y_t^4) = \sigma_e^4\left[3\phi^4 + 9\sigma_h^4 + 3 + 12\frac{2\pi}{\pi}(2\phi\sigma_b + 6\sigma_h^2 + 6\phi^2 + 12\phi^3)\frac{2\pi}{\pi} + 18\phi^2\sigma_h^2 + 24\frac{2\pi}{\pi}(2\phi\sigma_h^2)\right] \]

\[ = \sigma_e^4\left[3(\phi^4 + 3\sigma_b^4 + 1) + 12\frac{2\pi}{\pi}(1 + 2\sigma_b^2 + \phi^2)\phi\sigma_b + 6(\sigma_b^2 + \phi^2 + 3\phi^2\sigma_b^2)\right] \]

Thus,

\[ K(v) = \frac{E(y_t - \mu)^4}{(\text{Var}(y_t))^2} = \frac{3(\phi^4 + 3\sigma^4 + 1) + 12\frac{2\pi}{\pi}(1 + 2\sigma^2 + \phi^2)\phi\sigma_b + 6(\sigma^2 + \phi^2 + 3\phi^2\sigma^2)}{\left(1 + \phi^2 + 2\phi\sigma_b\frac{2\pi}{\pi} + \sigma_b^2\right)^2} \]

The autocorrelation function is given by

\[ \rho_0 = 1, \quad \rho_k = \frac{\phi + \frac{2\pi}{\pi}\sigma_h^2}{1 + \phi^2 + 2\phi\sigma_h^2\frac{2\pi}{\pi} + \sigma_h^2}, \quad \rho_k = 0 \quad k = 2, 3, 4, \ldots \]

**Theorem 2.4.** Let \( \{y_t\} \) be a Sign RCA-GARCH (1,1) time series satisfying conditions (i) and (ii) given by

\[ y_t = (\phi + |b_t| + \Phi s_t) y_{t-1} + \varepsilon_t \quad \text{(2.11)} \]

where

\[ \varepsilon_t = \sqrt{h_t} Z_t, \quad \text{(2.12)} \]

\[ h_t = \omega + \sum_{i=1}^{p} \alpha_i \varepsilon_{t-i}^2 + \sum_{j=1}^{q} \beta_j h_{t-j} \quad \text{(2.13)} \]

where \( Z_t \) and \( b_t \) are sequences of independent, identically distributed random variables with zeromean, variance given by \( \sigma_z^2 \) and \( \sigma_b^2 \) respectively, \( \omega, \alpha_i, \beta_j \) and \( \Phi \) are real parameters, satisfying the following conditions, \( \omega > 0, \alpha_i \geq 0, \beta_j \geq 0, |\Phi| \leq \omega \). Note: \( E(s_t^4) = 1 \), and in order to calculate the kurtosis, we observe that \( E(s_t^4) = 1 \). Then, we have the following moment properties

\[ E(y_t^2) = \frac{\omega \sigma_z^2}{\left[1 - \left(\phi^2 + 2\frac{2\pi}{\pi}\phi\sigma_b + \sigma_b^2 + \Phi^2\right)\left(1 - (\omega\sigma_z^2 + \beta)\right)\right]} \]

\[ E(y_{t-1}^4) = \frac{E(Z_t^4)}{\left(1 - \Phi^4 - \phi^4 - 6\Phi^2(\phi^2 + \sigma_b^2) - \frac{2\pi}{\pi}(12\Phi^2\sigma_b + 4\phi^2\sigma_b + 8\sigma_b^2) - 3\sigma_b^2(2\phi^2 + \sigma_b^2)\right)} \]

\[ \times \frac{6\sigma_z^4\left(2\phi\sigma_b + \frac{2\pi}{\pi}(2\phi\sigma_b + \sigma_b^2)\right)^2}{\left(1 - \frac{2\pi}{\pi}(2\phi\sigma_b + \sigma_b^2 + \Phi^2)\right)^2} \]

\[ \times \frac{E(h_t^2)}{\left(1 - \Phi^4 - \phi^4 - 6\Phi^2(\phi^2 + \sigma_b^2) - \frac{2\pi}{\pi}(12\Phi^2\sigma_b + 4\phi^2\sigma_b + 8\sigma_b^2) - 3\sigma_b^2(2\phi^2 + \sigma_b^2)\right)} \]

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\[
K^{(e)} = \frac{E(Z_i)}{\sigma_z^2 \left(1 - \Phi^2 - \phi^2 - 6\Phi^2 \left(\phi^2 + \sigma_b^2\right) - \frac{2}{\sqrt{\pi}} \phi \left(12\sigma_b^2 + 8\sigma_b^3\right) - 3\sigma_b^2 \left(2\phi^2 + \sigma_b^2\right)\right)}
\]

\[
E(y_i^2) = E\left((\phi + h_i + \Phi_s_i) y_{i-1} + \epsilon_i\right)
\]

\[
E\left(h_i \sigma_z^2 \right) = \frac{E(Z_i)}{1 - \left(\phi^2 + 2\frac{2}{\sqrt{\pi}} \phi \sigma_b + \sigma_b^2 + \Phi^2\right)}
\]

\[
E\left(y_{i-1}^4\right) = \frac{E(Z_i)}{1 - \Phi^4 - \phi^4 - 6\Phi^4 \left(\phi^2 + \sigma_b^2\right) - \frac{2}{\sqrt{\pi}} \phi \left(12\sigma_b^2 + 4\phi^2 \sigma_b + 8\sigma_b^3\right) - 3\sigma_b^2 \left(2\phi^2 + \sigma_b^2\right)}
\]

Proof:

\[
y_i = (\phi + h_i + \Phi_s_i) y_{i-1} + \epsilon_i
\]

\[
E\left(y_i^2\right) = E\left((\phi + h_i + \Phi_s_i) y_{i-1} + \epsilon_i\right)^2
\]

\[
E\left(y_{i-1}^4\right) = AE\left(h_i^2\right) + BE\left(h_i^2\right)^2
\]

where

\[
A = \frac{E(Z_i)}{\left(1 - 6\Phi^2 \left(\phi^2 + \sigma_b^2\right) - 12\phi \Phi \sigma_b^2 \frac{2}{\sqrt{\pi}} - 4\phi^3 \sigma_b \frac{2}{\sqrt{\pi}} - 6\phi^2 \sigma_b^2 - 8\phi^2 \sigma_b^3 - 8\phi \sigma_b^4 - 3\sigma_b^4 - \Phi^4 - \phi^4\right)}
\]

\[
B = \frac{6\sigma_z^4 \left(2\phi \sigma_b \frac{2}{\sqrt{\pi}} + \Phi^2 + \phi^2 + \sigma_b^2\right) \left(1 - \phi^2 + \Phi^2 + \sigma_b^2 + \Phi^2\right) \left(2\phi^2 + \sigma_b^2\right)}{\left(1 - \Phi^4 - \phi^4 - 6\Phi^4 \left(\phi^2 + \sigma_b^2\right) - \frac{2}{\sqrt{\pi}} \phi \left(12\sigma_b^2 + 4\phi^2 \sigma_b + 8\sigma_b^3\right) - 3\sigma_b^2 \left(2\phi^2 + \sigma_b^2\right)\right)}
\]
Using the facts that

\[
E\left( \frac{h^2}{j} \right) = \frac{1}{E\left( Z^4 \right) - \left( E\left( Z^4 \right) - 1 \right) \sum_j \psi_j^2}
\]

Thus, we have the following expression for the Kurtosis of the process.

\[
K^{(v)} = \frac{E\left( Z^4 \right)}{E\left( h^2 \right)} = \frac{1}{E\left( Z^4 \right) - \left( E\left( Z^4 \right) - 1 \right) \sum_j \psi_j^2}
\]
Note, that when $\sigma_b^2 = 0, \Phi = 0$, and $Z_t \sim N(0,1)$ in (2.15), the kurtosis of the process converges to
\[
K^{(v)} = 3 \left(1 - \left(\phi^2 + 2\sqrt{\frac{2}{\pi}} \phi \sigma_b + \sigma_b^2 + \Phi^2\right)\right) \left[1 - \left(\alpha + \beta\right)^2\right] \left[1 - \left(\alpha + \beta\right)^2 - 2\alpha^2\right]
\]
(2.15)

When $\phi = 0$ in (2.16), the kurtosis of the process converge to
\[
K^{(v)} = \frac{3\left(1 - \left(\alpha + \beta\right)^2\right)}{1 - \left(\alpha + \beta\right)^2 - 2\alpha^2} > 3
\]

**Theorem 2.5.** Suppose $y_t$ is a modified RCA model with GARCH $(p, q)$ innovations of the form
\[
y_t = (\phi + \beta) y_{t-1} + \varepsilon_t
\]
\[
\varepsilon_t = \sqrt{h_t} Z_t
\]
where $\varepsilon_t$ is an uncorrelated noise process with zero mean and with variance $\sigma_b^2$, and $Z_t$ is an uncorrelated noise process with zero mean and with variance $\sigma_z^2$. Then, we have the following relationship

\[
E\left(y_t^2\right) = \left[\frac{\sigma_z^2}{1 - \phi^2 - \frac{2}{\pi} \phi \sigma_b - \sigma_b^2}\right] E\left(h_t\right)
\]
(2.17)

\[
E\left(y_t^4\right) = \left[\frac{3\sigma_z^4}{1 - 4\phi^4 \sigma_b^2 - 6\phi^2 \sigma_b^2 - 8\phi^2 \sigma_b^2 - 8\phi \sigma_b^2 - 3\sigma_b^4}\right] E\left(h_t^2\right)
\]
(2.18)

\[
K^{(v)} = \left[\frac{3 - \phi^2 - \frac{2}{\pi} \phi \sigma_b - \sigma_b^2}{1 - 4\phi^4 \sigma_b^2 - 6\phi^2 \sigma_b^2 - 8\phi \sigma_b^2 - 3\sigma_b^4}\right] \left[E\left(Z_t^4\right) - \left(E\left(Z_t^4\right) - 1\right) \sum_{j=1}^{\infty} \psi_j^2\right]
\]
(2.19)
Proof: Let, \( u_i = \varepsilon_i^2 - h_i \)

\[
\varepsilon_i^2 - u_i = \omega + \sum_{j=1}^{q} \alpha_j \varepsilon_j^2 + \sum_{j=1}^{q} \beta_j h_{ij} \left[ 1 - \left( \sum_{j=1}^{q} \alpha_j B' + \sum_{j=1}^{q} \beta_j B' \right) \right] \varepsilon_i^2 = \omega + \left( 1 - \sum_{j=1}^{q} \beta_j B' \right) u_i 
\]

\[
\phi(B)\varepsilon_i^2 = \omega + \beta(B) u_i, \quad \phi(B) = 1 - \sum_{j=1}^{q} \phi_j B', \quad \phi_i = (\alpha_i + \beta_i) \]

\[
\beta(B) = \left( 1 - \sum_{j=1}^{q} \beta_j B' \right), \quad r = \max(p,q) 
\]

\[
E(h_i^2) = \frac{1}{E(Z_i^4) - (E(Z_i^4) - 1) \sum_{j=1}^{q} \psi_j^2}, \quad K^{(c)} = \frac{E(Z_i^4)}{E(Z_i^4) - (E(Z_i^4) - 1) \sum_{j=1}^{q} \psi_j^2} 
\]

Now we have

\[
E(y_i^2) - \phi^2 E(y_{i-1}^2) - 2 \sqrt{\frac{2}{\pi}} \phi E(y_{i-1}^2) \sigma_h - E(y_{i-1}^2) \sigma_h^2 = E(h_i) \sigma_h^2, 
\]

Thus,

\[
E(y_i^2) - \phi^2 E(y_{i-1}^2) - 2 \sqrt{\frac{2}{\pi}} \phi E(y_{i-1}^2) \sigma_h - E(y_{i-1}^2) \sigma_h^2 = E(h_i) \sigma_h^2, 
\]

and

\[
\left( 1 - \phi^2 - 2 \sqrt{\frac{2}{\pi}} \phi \sigma_h - \sigma_h^2 \right) E(y_i^2) = E(h_i) \sigma_h^2 
\]

\[
E(y_i^2) = \left[ \frac{\sigma_h^2 \left( 1 - \phi^2 - 2 \sqrt{\frac{2}{\pi}} \phi \sigma_h - \sigma_h^2 \right)}{E(h_i)} \right] 
\]

\[
E(y_i^2) = \frac{E(h_i^2) + 12 \phi E(y_{i-1}^2) h_i X_i^2 + 6 \phi E(y_{i-1}^2) h_i Z_i^2 + 6 \phi^2 E(y_{i-1}^2) h_i Z_i^2}{\left( 1 - 4 \phi^2 E[y_{i-1}^2] - 6 \phi^2 E[y_{i-1}^2] - 4 \phi^2 E[y_{i-1}^2] - 4 \phi^2 E[y_{i-1}^2] - 4 \phi^2 E[y_{i-1}^2] \right)} 
\]

\[
= \frac{3 \sigma_h^2 E(h_i^2) + 12 \phi \sigma_h \sigma_i^2 \left( \frac{2}{\pi} \right) E(y_{i-1}^2) E(h_i) + 6 \phi^2 \sigma_h \sigma_i^2 E(y_{i-1}^2) E(h_i) + 6 \phi^2 \sigma_h \sigma_i^2 E(y_{i-1}^2) E(h_i) + 6 \phi^2 \sigma_h \sigma_i^2 E(y_{i-1}^2) E(h_i)}{\left( 1 - 4 \phi^2 \sigma_h \left( \frac{2}{\pi} \right) - 6 \phi^2 \sigma_h^2 - 8 \phi^2 \left( \frac{2}{\pi} \right) - 8 \phi^2 \sigma_h - 6 \phi^2 \sigma_h^4 \right)} 
\]

\[
= \left[ \frac{3 \sigma_h^4 \left( 1 - 4 \phi^2 \sigma_h \left( \frac{2}{\pi} \right) - 6 \phi^2 \sigma_h^2 - 8 \phi^2 \left( \frac{2}{\pi} \right) - 8 \phi^2 \sigma_h - 6 \phi^2 \sigma_h^4 \right)}{E(h_i^2)} \right] \]

\[
+ \left[ \frac{12 \phi \sigma_h \sigma_i^2 \left( \frac{2}{\pi} \right) + 6 \phi^2 \sigma_h \sigma_i^2 + 6 \phi^2 \sigma_h^2}{\left( 1 - 4 \phi^2 \sigma_h \left( \frac{2}{\pi} \right) - 6 \phi^2 \sigma_h^2 - 8 \phi^2 \left( \frac{2}{\pi} \right) - 8 \phi^2 \sigma_h - 6 \phi^2 \sigma_h^4 \right)} \right] E(y_{i-1}^2) E(h_i) 
\]
Using the fact that 

\[ E\left(y_i^2\right) = \frac{\sigma^2_Z}{\left(1 - \phi^2 - \sqrt{2\phi\sigma_b - \sigma_b^2}\right)} E\left(h_i^2\right) \]

we have

\[ E(y_i^4) = \frac{3\sigma^4_Z}{\left(1 - 4\phi^4\sigma_b^2 - 8\phi^2\sigma_b^2 - 3\sigma_b^4\right)} E(h_i^2) + \frac{\left(12\phi\sigma^2_Z\sigma_b^2 + 6\sigma_b^2\sigma^2_Z + 6\phi^2\sigma^2_Z\sigma_b^2\right)}{\left(1 - 4\phi^4\sigma_b^2 - 8\phi^2\sigma_b^2 - 3\sigma_b^4\right)} \frac{\sigma^2_Z}{\left(1 - \phi^2 - \sqrt{2\phi\sigma_b - \sigma_b^2}\right)} E\left(h_i^2\right)^2 \]

For convenience let

\[ E\left(y_i^4\right) = AE\left(h_i^2\right) + BE\left(h_i^2\right)^2 \]

where,

\[ A = \frac{3\sigma^2_Z}{\left(1 - 4\phi^4\sigma_b^2 - 8\phi^2\sigma_b^2 - 3\sigma_b^4\right)} \]

\[ B = \frac{\left(12\phi\sigma^2_Z\sigma_b^2 + 6\sigma_b^2\sigma^2_Z + 6\phi^2\sigma^2_Z\sigma_b^2\right)}{\left(1 - 4\phi^4\sigma_b^2 - 8\phi^2\sigma_b^2 - 3\sigma_b^4\right)} \frac{\sigma^2_Z}{\left(1 - \phi^2 - \sqrt{2\phi\sigma_b - \sigma_b^2}\right)} \]

The Kurtosis of the process is given by

\[ K^{(y)} = \frac{E\left(y_i^4\right)}{E\left(y_i^2\right)^2} = \frac{AE\left(h_i^2\right) + BE\left(h_i^2\right)^2}{E\left(y_i^2\right)^2} \]

\[ = \left[AE\left(h_i^2\right) + BE\left(h_i^2\right)^2\right] \left(1 - \phi^2 - \frac{2\phi\sigma_b - \sigma_b^2}{\sigma^2_Z} E\left(h_i^2\right)^2\right) \]

\[ = A \left(1 - \phi^2 - \frac{2\phi\sigma_b - \sigma_b^2}{\sigma^2_Z} \right) E\left(h_i^2\right)^2 + B \left(1 - \phi^2 - \frac{2\phi\sigma_b - \sigma_b^2}{\sigma^2_Z} \right) \frac{\sigma^2_Z}{E\left(h_i^2\right)^2} \]

Using the fact that

\[ \frac{E\left(h_i^2\right)}{E\left(h_i^2\right)^2} = \frac{1}{E\left(Z_i^2\right) - E\left(Z_i^4\right) - 1} \sum_{j=1}^{\infty} \psi_j \]

we have...
2.2. Quadratic GARCH Model

Besides having excess kurtosis market returns may display seriously skewed distributions. Linear GARCH models cannot cope with such skewness, and therefore we can expect forecast of linear GARCH model to be biased for skewed time series. To deal with this problem non-linear GARCH models are introduced, which take into account skewed distributions. The QGARCH model differs from the classical GARCH model by Theorem 2.6.

Consider the general class of RCA QGARCH (1,1) Volatility Models for the time series $y_t$, where

\[
y_t = \sqrt{h_t} Z_t
\]

This model reduces to the GARCH (1,1) model when the shift parameters $\delta_3 = 0$. The QGARCH model can improve upon the standard GARCH since they can cope with positive (or negative) skewness.

Theorem 2.6. Consider the general class of RCA QGARCH (1,1) Volatility Models for the time series $y_t$, where

\[
y_t = \sqrt{h_t} Z_t
\]

where $Z_t \sim N(0, \sigma^2_t)$ and $\alpha_t \sim N(0, \sigma^2_{\alpha_t})$. Then, we have the following moment properties
Proof: is easy and is omitted.

**Theorem 2.7.** Consider the special class of RCA QGARCH (1,1) Sign Volatility Models for the time series $y_t$, where

$$
y_t = (\phi + |b_t|)y_{t-1} + \varepsilon_t
$$

$$
e_t = \sqrt{h_t} Z_t
$$

$$
h_t = (\delta_0 + \delta_2 \varepsilon_{t-1}^2 + \delta_1 y_{t-1} + 2\delta_1 \varepsilon_{t-1} y_{t-1})
\quad + (\delta_2 + [a_{t-1} + \Phi s_{t-1}]) y_{t-1}^2
$$

where $Z_t \sim N(0, \sigma_Z^2)$, $a_t \sim N(0, \sigma_a^2)$, at $\sim N(0, 2\alpha a)$ and $b_t \sim N(0, \sigma_b^2)$ are sequences of independent, identically-distributed random variables with zero mean, variance given by $\sigma_Z^2$ and $\sigma_a^2$ and $\sigma_b^2$ respectively, and

$$
s_t = \begin{cases} 
+1 & \text{if } y_t > 0 \\
0 & \text{if } y_t = 0 \\
-1 & \text{if } y_t < 0
\end{cases}
$$

Note: $E(s_t^2) = 1$, and in order to calculate the kurtosis, we observe that $E(s_t^4) = 1$. Then, we have the following moment properties.

$$
E(h_t) = \frac{\left(1 - \left(\phi^2 + 2\phi\sigma_b \frac{2}{\sqrt{\pi}} + \sigma_b^2\right)\right)^2}{(1 - \delta_0)}
$$

$$
E(y_t^2) = \frac{\sigma_Z^2 \left(1 - \left(\phi^2 + 2\phi\sigma_b \frac{2}{\sqrt{\pi}} + \sigma_b^2\right)\right)^2}{(1 - \delta_0)}
$$

$$
E(y_t^4) = \frac{3\sigma_Z^4}{\left(1 - \left(\phi^4 + 3\sigma_b^4 + 4\phi^3 \sigma_b \frac{2}{\sqrt{\pi}} + 6\phi^2 \sigma_b^2 + 8\phi \frac{2}{\sqrt{\pi}} \sigma_b^2\right)\right)}
\left(1 - \left(\phi^2 + 2\phi\sigma_b \frac{2}{\sqrt{\pi}} + \sigma_b^2\right)\right)
$$

$$
K^{(y)} = \frac{\left(1 - \left(\phi^2 + 2\phi\sigma_b \frac{2}{\sqrt{\pi}} + \sigma_b^2\right)\right)^2}{\sigma_Z^4} A E(h_t) + B
$$

where

$$
A = \frac{3\sigma_Z^4}{\left(1 - \left(\phi^4 + 3\sigma_b^4 + 4\phi^3 \sigma_b \frac{2}{\sqrt{\pi}} + 6\phi^2 \sigma_b^2 + 8\phi \frac{2}{\sqrt{\pi}} \sigma_b^2\right)\right)}
$$

$$
B = \frac{\sigma_Z^4 \left(6\sigma_b^4 + 6\phi^2 + 12\phi \sigma_b \frac{2}{\sqrt{\pi}} \right) \left(1 - \left(\phi^2 + 2\phi\sigma_b \frac{2}{\sqrt{\pi}} + \sigma_b^2\right)\right)^{-1}}{\left(1 - \left(\phi^4 + 3\sigma_b^4 + 4\phi^3 \sigma_b \frac{2}{\sqrt{\pi}} + 6\phi^2 \sigma_b^2 + 8\phi \frac{2}{\sqrt{\pi}} \sigma_b^2\right)\right)}
$$
Proof: is easy and is omitted.

3. Option Pricing with Volatility

Option pricing based on the Black-Scholes model is widely used in the financial community. The BlackScholes formula is used for the pricing of European-style options. The model has traditionally assumed that the volatility of returns is constant. However, several studies have shown that asset returns exhibit variances that change over time [9,10,] and others derived closed form option pricing formulas for different models which are assumed to follow a GARCH volatility process. Most recently, Gong et al. [11] derive an expression for the call price as an expectation with respect to random GARCH volatility. The model is then evaluated in terms of the moments of the volatility process. Their results indicate that the suggested model outperforms the classic Black-Scholes formula. Here we apply [11] and propose an option pricing model with RCA GARCH volatility as follows:

\[ E(h^n) = \frac{M_1 + M_2}{(1-\delta^n) - M_n A} + \frac{(M_1 + M_4 + M_5 + M_6 + M_7 + M_8)\sigma^2}{(1-\delta^n) - M_n A} E[h^n] + \frac{M_n B (E[h^n])^2}{(1-\delta^n) - M_n A} \] (2.35)

\[ M_1 = (\delta^n_b + 2\delta^n_0 \delta^n_1 + \delta^n_2) \]

\[ M_2 = \left(2\delta^n_1 \delta^n_2 \left(1 - \left(\phi^2 + 2\phi \sigma_c \frac{\sigma}{\pi} + \sigma_c^2\right)\right)\right) (\delta^n_0 + \delta^n_2 \delta^n_1) \]

\[ M_3 = 2\delta^n_0 \delta^n_2, \quad M_4 = 6\delta^n_0^2 \delta^n_1^2 \]

\[ M_5 = \left(1 - \delta^n_1 \left(1 - \left(\phi^2 + 2\phi \sigma_c \frac{\sigma}{\pi} + \sigma_c^2\right)\right)\right) (\delta^n_0 + \delta^n_2 \delta^n_1) \]

\[ M_6 = 2\delta^n_0 \sigma_c \frac{\sigma}{\pi} \]

\[ M_7 = \left(1 - \delta^n_1 \left(1 - \left(\phi^2 + 2\phi \sigma_c \frac{\sigma}{\pi} + \sigma_c^2\right)\right)\right) \sigma_c^2 (\delta^n_0 + \delta^n_2 \delta^n_1) \]

\[ M_8 = 2\delta^n_0 \delta^n_2 \sigma_c \frac{\sigma}{\pi} \]

\[ M_9 = \left(\Phi^2 + 2\sigma_c \frac{\sigma}{\pi} + \sigma_c^2\right) \]

\[ \frac{dS_t}{S_t} = \sigma_{S_t} dW_t \] (3.1)

\[ y_t = \log \left(\frac{S_t}{S_{t-1}}\right) - E \left[ \log \left(\frac{S_t}{S_{t-1}}\right) \right] = \sigma y_t \] (3.2)

\[ \Phi(B) \sigma^2 = \omega + \beta(B) y_t^2 \] (3.3)

where \( S_t \) is the price of the stock, \( r \) is the risk-free interest rate, \( \{W_t\} \) is a standard Brownian motion, \( \sigma_t \) is the time-varying RCA GARCH volatility process, \( \{Z_t\} \) is a sequence of i.i.d. random variables with zero mean and unit variance and \( \Phi(B) \) and \( \beta(B) \) have been defined in (1.5). The price of a call option can be calculated using the option pricing formula given in [11]. The call price is derived as a first conditional moment of a truncated lognormal distribution under the martingale measure, and it is based on estimates of the moments of the GARCH process. The call price based on the Black-Scholes model with seasonal GARCH volatility is given by:
\[ C(S,T) = S\left[ f[E(\sigma_i^2)] + \frac{1}{2} f''[E(\sigma_i^2)] \left( \frac{1}{3} \kappa^{(i)} - 1 \right) E^2(\sigma_i^2) \right] \]
\[ - Ke^{-\tau} \left[ g[E(\sigma_i^2)] + \frac{1}{2} g''[E(\sigma_i^2)] \left( \frac{1}{3} \kappa^{(i)} - 1 \right) E^2(\sigma_i^2) \right], \]  
(3.4)

where \( f \) and \( g \) are twice differentiable functions, \( S \) is the initial value of \( S_t \), \( K \) is the strike price, \( T \) is the expiry date, \( \sigma_t \) is a stationary process with finite fourth moment, and \( \kappa^{(i)} = \frac{E(y_i^2)}{[E(y_i^2)]^2} \).

Also,

\[ f[E(\sigma_i^2)] = N(d) = N \left( \frac{\log(S/K) + rT + \frac{1}{2} E(\sigma_i^2)}{\sqrt{E(\sigma_i^2)}} \right), \]
\[ g[E(\sigma_i^2)] = N(d - \sqrt{E\sigma_i^2}) = N \left( \frac{\log(S/K) + rT - \frac{1}{2} E(\sigma_i^2)}{\sqrt{E\sigma_i^2}} \right), \]

and

\[ f''[E(\sigma_i^2)] = \frac{1}{2\sqrt{\pi}} \left[ \frac{E(\sigma_i^2) - 2(\log(S/K) + rT)}{4E(\sigma_i^2)\sqrt{E(\sigma_i^2)}} \right] \left[ \frac{E(\sigma_i^2)^2 - 4(\log(S/K) + rT)^2}{8E(\sigma_i^2)^2} \right] \]
\[ + \frac{6(\log(S/K) + rT) - E(\sigma_i^2)}{8E(\sigma_i^2)^2} \exp \left\{- \frac{2(\log(S/K) + rT + E(\sigma_i^2))^2}{8E(\sigma_i^2)} \right\}, \]

\[ g''[E(\sigma_i^2)] = \frac{1}{2\sqrt{\pi}} \left[ \frac{E(\sigma_i^2) + 2(\log(S/K) + rT)}{4E(\sigma_i^2)\sqrt{E(\sigma_i^2)}} \right] \left[ \frac{E(\sigma_i^2)^2 - 4(\log(S/K) + rT)^2}{8E(\sigma_i^2)^2} \right] \]
\[ + \frac{6(\log(S/K) + rT) + E(\sigma_i^2)}{8E(\sigma_i^2)^2} \exp \left\{- \frac{2(\log(S/K) + rT - E(\sigma_i^2))^2}{8E(\sigma_i^2)} \right\}, \]

where \( N \) denotes the standard normal CDF, and under the option pricing model with RCA GARCH volatility,

\[ E(\sigma_i^2) = \frac{\omega}{\left( 1 - \sum_{i=1}^p \phi_i \right) \left( 1 - \sum_{i=1}^p \Phi_i \right)}, \quad \kappa^{(i)} = \frac{3}{3 - 2\sum_{j=1}^p \Psi_j}. \]

4. Concluding Remarks

Financial time series exhibit excess kurtosis and in this paper, we propose various classes of RCAGARCH volatility models and derive the kurtosis in terms of model parameters. We consider time series models such as RCA with GARCH errors and quadratic GARCH errors. The models introduced here extend and complement the existing volatility models in the literature to RCA models with quadratic GARCH models by introducing more general structures. The results are primarily oriented to financial time series applications. Financial time series often meet the large data set demands of the volatility models studied here. Also, financial data dynamics and higher order moments are of interest to many market participants. Specifically, we consider the Black-Scholes model with RCA GARCH volatility and show that these moments can be used to evaluate the call price for European options.
REFERENCES


