Extending Multi-Period Pluto and Tasche PD Calibration Model Using Mode LRDF Approach

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Abstract
The intention of this paper is to propose extension to the Pluto and Tasche PD calibration model for low default portfolios that could produce more stable LRDF estimates and eliminate the necessity of quartile choice, while preserving adequate level of conservatism. Multi-period Pluto and Tasche model allows us to fulfill Basel committee requirements regarding long-term LRDF calibration even for portfolios with no observable defaults. The main drawback of that approach is a very strict requirement for the sample: only borrowers that are observable to the bank within each point on long-term horizon could be used as observations. Information regarding rating migrations, borrowers that arrived in the portfolio after sample cutoff date and borrowers that left the portfolio before the end of long-term calibration horizon should be excluded from the sample. Proposed Mode approach pairs Pluto and Tasche model with mode LRDF estimator (proposed by Canadian OSFI), as the results, it eliminates drawbacks of the original Pluto and Tasche model.

Keywords
Credit Risk, Low Default Portfolio, PD Calibration

1. Introduction
This paper is dedicated to ultra-low default portfolios. Ultra-low default portfolio is the portfolio for which we haven’t got enough historical defaults to estimate discriminatory power of the ranking model (e.g. calculate Gini coefficient) and/or central tendency. Therefore it’s impossible to perform classical central tendency probability of default (PD) calibration using one of the common approaches described, for example, in Tasche (2009) [1]. Nevertheless, using Pluto and Tasche (2005) [2] model it is possible to statistically calibrate PD for ultra-low
default portfolios even in case of absence of historical defaults.

The idea of Pluto and Tasche approach is the following (details see in Pluto K. and Tasche D. (2005) [2]). We assume that PD assigned to rating classes (still to be estimated) reflects the decreasing credit-worthiness of the grades, i.e. PD is a non-decreasing function of rating class:

\[ pd_A \leq pd_B \leq pd_C. \] (1)

The most prudent assumption of PD for any given rating is its equality to the PD of the worst rating within the rating grade. Therefore, we assume \( pd_A = pd_B = pd_C \) for rating class A, \( pd_B = pd_C \) for B rating and so forth. Assuming unconditional independence of the default events, the probability of observing no defaults turns out to be \( (1 - pd_A)^{n_A} \) where \( n_i \) denotes number of borrowers in the rating class \( i \). In order to find PD estimate for rating class A, we should solve the following inequality (given some confidence level \( \gamma \) for PD estimation):

\[ 1 - \gamma \leq (1 - pd_A)^{n_A + n_B + n_C}. \] (2)

In case of available \( \kappa = \kappa_A + \kappa_B + \kappa_C \) (number of defaults should be less or equal to the number of borrowers in each rating class: \( k_i \leq n_i, \forall i \) ) historical defaults we simply replace probability estimate by the CDF (cumulative distribution function) of binomial distribution:

\[ 1 - \gamma \leq \sum_{i=0}^{\kappa} \binom{n_A + n_B + n_C}{i} pd_A^i (1 - pd_A)^{n_A + n_B + n_C - i}. \] (3)

Using the same logic we could estimate \( pd_B, pd_C \):

\[ 1 - \gamma \leq \sum_{i=0}^{\kappa_B} \binom{n_B + n_C}{i} pd_B^i (1 - pd_B)^{n_B + n_C - i}. \] (4)

\[ 1 - \gamma \leq \sum_{i=0}^{\kappa_C} \binom{n_C}{i} pd_C^i (1 - pd_C)^{n_C - i}. \] (5)

Generalized rule for PD calibration that is used under original P&T model that could be expressed as:

\[ 1 - \gamma \leq P[\text{No more than } k \text{ defaults observed}] \] (6)

As the result, solving (6) for each separately leads us to original P&T one-period PD calibration approach that obviously don’t comply with Basel committee requirement for long-term LRPD estimation period [3]. In order to comply with the requirements, one can use pooling, e.g. replacing the one-year number of borrowers in a grade with the sum of the borrower numbers of this grade over the years (analogously for the numbers of defaulted borrowers). However, when turning to the case of cross-sectionally and intertemporally correlated default events, pooling does not allow for an adequate modelling—there is no way to incorporate into the model information about correlation between borrowers and autocorrelation of the systemic factor. Other option is to use multi-period extension of the Pluto and Tasche (P&T) model which takes into account correlation between borrowers and autocorrelation of the systemic factor.

Multi-period P&T model assumes (as in Vasicek, Oldrich A. (1997) [4]) that a default of borrower is triggered if the change in a value of its assets results in a value lower than some default threshold \( c \):

\[ V_{i,t} = \sqrt{\rho S_t} + \sqrt{1 - \rho} \xi_{i,t} \] (7)

where \( \rho \) stands for the so-called asset correlation, \( S_t \) is the realization of the systematic factor in year \( t \), and \( \xi_{i,t} \) denotes the idiosyncratic (or borrower-specific) component of the change in value. The cross-sectional dependence of the default events stems from the presence of the systematic factor \( S_t \). Both systematic and idiosyncratic factors are standard normally distributed, and idiosyncratic factors are i.i.d., while joint distribution of \( S_t \) is completely determined by the correlation matrix \( RC \) (which could be expressed using time-distance function).

Borrower \( i \)'s default occurs in year \( t \) if:
where default threshold \( c \) could be calibrated from unconditional PD:

\[
c = \Phi^{-1}(pd)
\]

with \( \Phi \) denoting the standard normal distribution function.

Following Bluhm, C., Overbeck, L. and C. Wagner (2003) [5], conditional on particular systematic factor realization default probability in a given year could be expressed as the following link function:

\[
G(pd, \varrho, S_t) = \Phi \left( \frac{\Phi^{-1}(pd) - \sqrt{\rho}S_t}{\sqrt{1-\rho}} \right)
\]

(10)

From (7), (8) and by using of the conditional independence of the \( V_{i,1}, \cdots, V_{i,T} \) and given systematic factors \( S_1, \cdots, S_T \), we obtain:

\[
\pi(S_1, \cdots, S_T) = P\left[ \text{Borrower } i \text{ defaults} \mid S_1, \cdots, S_T \right]
\]

\[
= P\left[ \min_{t=1, \cdots, T} V_{i,t} \leq \Phi^{-1}(pd) \mid S_1, \cdots, S_T \right]
\]

\[
= 1 - \prod_{t=1}^{T} \left( 1 - G(pd, \varrho, S_t) \right)
\]

(11)

Right hand side of (6) could be written as:

\[
P[\text{No more than } k \text{ defaults observed}]
\]

\[
= \sum_{l=0}^{k} \left( \sum_{l=0}^{N} \left[ \pi(S_1, \cdots, S_T)^l \left( 1 - \pi(S_1, \cdots, S_T) \right)^{N-l} \right] \right)
\]

(12)

Solving inequality (6) in each rating class separately for PD values given Equation (12) leads us to PD estimates according to original multi-period Pluto & Tasche model.

Above described P&T model has obvious drawbacks in case of multi-period calibration. Pooling approach is unable to take into account correlation between borrowers and systematic factor time series, as the result, we get PD estimates that has zero sensitivity to really observable economic effect of default clusterization.

Multi-period P&T model considers only the borrowers that were present at the beginning of the observation period. Therefore we lose significant amount of information for stable portfolios and decreasing stability of PD estimates. Moreover, in case of shifts in the portfolio (like significant growth of the portfolio or high prepayment rate) the model is totally inapplicable.

The other drawback of P&T model is the necessity of the quartile \( \gamma \) choice. Mode \( \gamma = 50\% \) could be seen as not enough conservative value (could underestimate true LRDF in more than 50\% cases—see Monte Carlo simulation results below). Current practices of quartile choice relate to external factors such as external ratings, CDS spreads and so on. Those practices are suitable for quartile choice under ICAAP, but a PD value of the borrower is an objective figure that obviously should not vary between banks and depend on risk profile of a lender.

2. P&T Model Extension Using Mode Calibration

The idea of the proposed approach is to pair P&T model with the approach proposed by Canadian financial institutions regulator, the Office of the Superintendent of Financial Institutions (OSFI), as a possible way to increase conservatism of LRPD estimation. For detail illustrations, please refer to OSFI document: “Risk Quantification of IRB Systems at IRB Banks” [6] and Miu P. and Ozdemir B. (2008) [7].

The idea of the approach is to find the LRPD\(_{mod}\) given the average long-term default rate in a given rating class as:
\[ PD_{ave} = \text{mod} \left( G \left( pd_{mod}, \theta, S_i \right) \right) \]  

(13)

where \( \text{mod} ( \cdot ) \) is the mode of the distribution of the average default rate given the arguments of the link function (4). Since there is no explicit analytical Equation to relate the mode of the distribution with the underlying LRPD, this approach involves changing the assumed value of LRPD until we obtain a mode (through simulations) of the distribution of the simulated average default rate \( PD_{ave} \), which is identical (or close enough) to the observed average default rate average \( PD_{ave} \).

Proposed approach consists of the following steps (assuming we have historical data for \( T \) periods in \( R \) rating classes, for each period—\( n_{t,r} \) borrowers and \( k_{t,r} \) defaults):

1) Apply to each historical period \( t = 1, \cdots, T \) one-period P&T model (1) with the quartile \( \gamma = 50\% \). The output of the model is a set of one period PDs: \( PD_{t,r}^{P&T} \);

2) Let’s assume that \( PD_{t,r}^{P&T} \) is a true default rate within period \( t \) in rating class \( r \) (instead of observed default rate which could be distorted due to LDP nature of the portfolio);

3) Apply OSFI LRPD \( \text{mod} \) approach to \( PD_{t,r}^{P&T} \).

a) Calculate mean of P&T default rates in each rating class:

\[ \bar{PD}_{r}^{P&T} = \frac{\sum_{t=1}^{T} PD_{t,r}^{P&T}}{T} \]  

(14)

b) Simulate \( S \) using normal multivariate distribution with correlation matrix \( RC \) for \( n = 1, \cdots, N \) scenarios.

c) For each rating class \( R \) find such \( PD_{r}^{\text{mod}} \) that satisfies Equation:

\[ \bar{PD}_{r}^{P&T} = \text{mod} \left( E \left[ G \left( PD_{r}^{\text{mod}}, \theta, S_t \right) \right] \right), \]  

(15)

where expectation is taken with respect to scenarios.

Equation (15) provides us with the values of \( PD_{r}^{\text{mod}} \) that correspond to claimed Mode LRDF calibration approach. In order to prove the efficiency of the proposed estimator we could use Monte-Carlo simulations.

3. Monte-Carlo Simulation Details

All examples are based on the portfolio with 3 rating classes A, B, C with the number of borrowers for the first portfolio: \( n_A = 100, n_B = 400, n_C = 300 \); second portfolio: \( n_A = 50, n_B = 200, n_C = 150 \). One-factor model cross-sectional correlation is constant \( \rho = 0.12 \), while intertemporal correlation, described by correlation matrix \( RC \), depends on time distance:

\[ r_{c,i,j} = \nu^{i-j}, \quad s, t = 1, \cdots, T \]  

(16)

where \( \nu = 0.3 \), number of periods \( T \) is 5 (as required by [3]).

In order to simulate the loss of information in multi-period P&T model random repayment events are incorporated into the simulations. Each borrower in each period could leave the portfolio due to debt repayment with the probability \( P \) [repayment]. Simulations were held with two different values: 15% and 30%. We assume that the bank is keen to hold stable rating structure, i.e. in case of default or repayment, the borrower is replaced by another one in the same rating class. Therefore in case of the Proposed Mode approach we have the same rating structure at the beginning of each period.

True LRDF rates (PD variables values in Equation (10)) by rating classes are: \( n_A = 0.001, n_B = 0.0025, n_C = 0.005 \).

1) Monte-Carlo simulation schema (applied for each combination of the repayment probability and portfolio) is the following:

a) Generate 1000 portfolio scenarios, in each scenario;

b) Generate \( S_t, t = 1, \cdots, 5 \) values using correlation matrix (16);

c) Test each borrower for default case (using its rating LRDF value and link function (10));

d) Simulate repayments with the uniform probability \( P \) [repayment];

i) Produce two types of tables (number of borrowers and defaults) by rating classes,
ii) Taking into account repayments, not taking into account default replacements (multi-period P&T model).

2) Under assumption of stable rating structure—all repaid and defaulted borrowers are replaces in the next period (pooled P&T model, Proposed Mode based approach):
   a) Estimate LRDF values using the following approaches;
   b) P&T with $\gamma = 50\%$;
   c) P&T with $\gamma = 90\%$;
   d) Proposed Mode based approach.
   e) Pooled P&T model.

3) For each approach calculate the following statistics:
   a) Mean of LRDF estimates;
   b) Standard deviation of LRDF estimates;
   c) Percentage of LRDF underestimations;
   d) Plot LRDF estimates density (using build-in «R» function «density» for kernel).

Simulation results for different combinations of prepayment rates and portfolios are provided in Tables 1-4. LRDF estimates density graphs are provided in the Figures 1-4.

<p>| Table 1. Portfolio: (50, 200, 150); Repayment probability = 15%. |</p>
<table>
<thead>
<tr>
<th>Rating LRDF</th>
<th>Avg PD</th>
<th>SD</th>
<th>LRDF underestimation rate</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>PT 0.5</td>
<td>PT 0.9</td>
<td>Mode</td>
</tr>
<tr>
<td>0.10%</td>
<td>0.36%</td>
<td>0.94%</td>
<td>0.58%</td>
</tr>
<tr>
<td>0.25%</td>
<td>0.40%</td>
<td>1.03%</td>
<td>0.64%</td>
</tr>
<tr>
<td>0.50%</td>
<td>0.62%</td>
<td>1.67%</td>
<td>1.08%</td>
</tr>
</tbody>
</table>

<p>| Table 2. Portfolio: (50, 200, 150); Repayment probability = 30%. |</p>
<table>
<thead>
<tr>
<th>Rating LRDF</th>
<th>Avg PD</th>
<th>SD</th>
<th>LRDF underestimation rate</th>
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<tbody>
<tr>
<td></td>
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<td></td>
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</tr>
<tr>
<td></td>
<td>PT 0.5</td>
<td>PT 0.9</td>
<td>Mode</td>
</tr>
<tr>
<td>0.10%</td>
<td>0.65%</td>
<td>1.66%</td>
<td>0.56%</td>
</tr>
<tr>
<td>0.25%</td>
<td>0.72%</td>
<td>1.83%</td>
<td>0.62%</td>
</tr>
<tr>
<td>0.50%</td>
<td>1.16%</td>
<td>3.08%</td>
<td>1.07%</td>
</tr>
</tbody>
</table>

<p>| Table 3. Portfolio: (100, 400, 300); Repayment probability = 15%. |</p>
<table>
<thead>
<tr>
<th>Rating LRDF</th>
<th>Avg PD</th>
<th>SD</th>
<th>LRDF underestimation rate</th>
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<tr>
<td></td>
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</tr>
<tr>
<td></td>
<td>PT 0.5</td>
<td>PT 0.9</td>
<td>Mode</td>
</tr>
<tr>
<td>0.10%</td>
<td>0.32%</td>
<td>0.78%</td>
<td>0.47%</td>
</tr>
<tr>
<td>0.25%</td>
<td>0.35%</td>
<td>0.85%</td>
<td>0.53%</td>
</tr>
<tr>
<td>0.50%</td>
<td>0.52%</td>
<td>1.29%</td>
<td>0.83%</td>
</tr>
</tbody>
</table>

<p>| Table 4. Portfolio: (100, 400, 300); Repayment probability = 30%. |</p>
<table>
<thead>
<tr>
<th>Rating LRDF</th>
<th>Avg PD</th>
<th>SD</th>
<th>LRDF underestimation rate</th>
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<tr>
<td></td>
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<tr>
<td></td>
<td>PT 0.5</td>
<td>PT 0.9</td>
<td>Mode</td>
</tr>
<tr>
<td>0.10%</td>
<td>0.59%</td>
<td>1.37%</td>
<td>0.49%</td>
</tr>
<tr>
<td>0.25%</td>
<td>0.65%</td>
<td>1.51%</td>
<td>0.54%</td>
</tr>
<tr>
<td>0.50%</td>
<td>0.97%</td>
<td>2.33%</td>
<td>0.84%</td>
</tr>
</tbody>
</table>
In all cases Proposed Mode approach is more conservative than multi-period P&T $\gamma = 50\%$ model and Pooling approach and less conservative than multi-period P&T $\gamma = 90\%$. The maximum LRDF underestimation rate for Mode approach is less than 30%, therefore the approach is enough conservative and there is no need in additional parameters involvement (in terms of max LRDF underestimation rate Mode approach is almost twice more conservative than P&T $\gamma = 50\%$ model and Pooling).

The standard deviations of the estimates is on par with multi-period P&T $\gamma = 50\%$ model and Pooling approach, at the same time Mode approach SD is significantly smaller than multi-period P&T $\gamma = 90\%$ SD. As the result, additional conservatism in the Mode approach does not lead to decreasing in the stability of the estimates.

Increase in repayment probability leads to shift in mean of PD estimates and increased SD for multi-period P&T models (despite there is no economic reason to increasing PD due to increase in repayment rates). Obviously, it’s due to significant sample restrictions. Proposed Mode approach has not got such restriction, as the result, both mean and standard deviations of estimates are stable despite repayment rate changes.

4. Conclusions
The Proposed Mode approach has the following advantages:
1) It takes into account all available information about borrowers and defaults in comparison to serious sample constrains of multi-period P&T model. As the result, estimates don’t depend on the repayment profile of the portfolio: increase in repayments does not lead to negative changes in standard deviation of the results and shifts in PD values.

2) It takes into account cross-sectional and intertemporal correlation structure (see steps 3. b and 3. c in Monte Carlo simulation schemas) that leads to more accurate and theoretically justified LRDF estimates.

3) It provide enough level of conservatism (the approach is a much more conservative approach than pooled and one-period P&T model $\gamma = 50\%$) while maintaining the same of standard deviation of estimates, therefore there is no need to use additional quartile parameter in order to increase conservatism.

4) Additional conservatism of the approach does not lead to increase in standard deviation values as it happens in case of increased quartile in P&T model.

References


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