Optimal Stopping Time to Buy an Asset When Growth Rate Is a Two-State Markov Chain

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Abstract

In this paper we consider the problem of determining the optimal time to buy an asset in a position of an uptrend or downtrend in the financial market and currency market as well as other markets. Asset price is modeled as a geometric Brownian motion with drift being a two-state Markov chain. Based on observations of asset prices, investors want to detect the change points of price trends as accurately as possible, so that they can make the decision to buy. Using filtering techniques and stochastic analysis, we will develop the optimal boundary at which investors implement their decisions when the posterior probability process reaches a certain threshold.

Keywords
Optimal Stopping Time, Posterior Probability, Threshold

1. Introduction

In [1], the authors consider the problem of determining the optimal time to sell a property while price growth rate is a random variable that takes the value of the given set. Under the assumptions of the problem considered in [1], growth rate only gets one of the possible values that do not change from this value to other values, which means that transition probability density is 0; but at a time I do not know the accuracy of price growth rate and the probability of receiving a certain value of growth rate also changes over time.

The authors in [2] also have much optimal stopping time in mathematical finance, but this is the classical problem and less common in practice because of its assumptions.

The authors in [3] consider the optimal stopping time problem when the growth rate of price process is not a random variable but in many cases.
In [4] the authors consider the problem to find the optimal time to sell when the price growth rate is Markov chain, however their approach is different from our method in this paper and the results are also different.

In this paper we consider the price process is described by geometric Brownian motion when its drift (price growth) is the Markov chain with two states 0 and 1: (0: decrease, downtrend; 1: increase, uptrend).

There is a phenomenon known as the momentum of the stock price, which means that the stock price has increased or is increasing, it will be increased in the future (usually near future), and a stock price has decreased or is decreasing, it will be decreased in the future. The investment is based on the momentum of price process that is called the momentum trading or trend following trading. By this way if an investor holding an asset wants to sell the asset, he will have to wait until a bull appeared and continues to hold such assets to the next price increase (in momentum) and when momentum is no longer available, i.e. prices going down or just starting going down, then he decides to sell. Similarly, if a person wants to buy an asset, he will wait for the appearance of an opportunity of going down in price and wait for prices to go down further (in momentum) until no longer falling price, then he decides to buy. According to this way of investment, investors are expected to buy the property at the bottom of market and sell the property at the highest point of the market.

In this paper, we analyze how to select buy and sell optimal strategies for a momentum trading. More precisely, we seek to maximize the expected profit from a momentum trade.

The method we use to study in this paper is the martingale theory, change of measure and the optimal stopping time is referred in the literature [2] [5] and [6].

2. Buying Asset Problem

Now we consider the case \( a, t \geq 0 \), is a Markov chain with two states 0 and 1 that \( P(a = a_0) = \pi_0, P(a = a_1) = 1 - \pi_0 \) and only capable of moving from state 0 to state 1 with transition density as follows \( Q = \begin{bmatrix} -\lambda & \lambda \\ 0 & 0 \end{bmatrix} (\lambda > 0) \)

At the time of \( t > 0 \) we put \( \pi_t = P(a = a_1 | \mathcal{F}_t) \) where \( \{\mathcal{F}_t\}_{0 \leq t} \) is the filter generated by \( X \).

We now consider the problem: Find \( \mathcal{F}_t \)-stopping time \( \tau, 0 \leq \tau \leq T \) such that:

\[
U = \inf_{0 \leq t \leq T} \mathbb{E}\left[e^{-\tau X_t}\right] \tag{2.1}
\]

According to the theorem 9.1 (view [5]), posterior probability process \( \pi_t \) satisfying:

\[
d\pi_t = \lambda (1 - \pi_t) dt + \pi_t \frac{a_h X_t - \left[a_h \pi_t + a_i (1 - \pi_t)\right] X_t}{\sigma X_t} d\mathcal{W}_t = \lambda (1 - \pi_t) dt + \pi_t \left(\sigma \pi_t - \sigma\pi_t\right) d\mathcal{W}_t\]

where

\[
d\mathcal{W}_t = \frac{dX_t - \left[(1 - \pi_t) a_i + \pi_t a_h\right] X_t dt}{\sigma X_t} = \left\{a_i - \left[(1 - \pi_t) a_i + \pi_t a_h\right]\right\} X_t dt + \sigma X_t dW_t\]

and \( (\mathcal{W}, \mathcal{F}_t) \) is a Brownian motion and the filter generated by \( \mathcal{W} \) coincides with filter \( \mathcal{F}_X \).

Thus the process \( X_t \) satisfies the equation:

\[
\frac{dX_t}{X_t} = \mathbb{E}\left[a_i | \mathcal{F}_X\right] dt + \sigma d\mathcal{W}_t = \left[(1 - \pi_t) a_i + \pi_t a_h\right] dt + \sigma d\mathcal{W}_t.
\]

Then the process \( X_t \) and the posterior probability process \( \pi_t \) satisfies the equation:
\[
\begin{align*}
\frac{dX_t}{X_t} &= (a_t + \pi_t (a_t - a_i)) dt + \sigma d\tilde{W}_t, \\
\frac{d\pi_t}{\pi_t} &= \lambda_t (1 - \pi_t) dt + \pi_t (1 - \pi_t) \left( \frac{a_h - a_l}{\sigma} \right) d\tilde{W}_t \\
\end{align*}
\]
(2.2)

Put $\Phi_t = \frac{\pi_t}{1 - \pi_t}$. According to Ito formula we have:

\[
\begin{align*}
d\Phi_t &= \frac{1}{(1 - \pi_t)^2} d\pi_t + \frac{1}{(1 - \pi_t)^3} \left[ \pi_t (1 - \pi_t) \left( \frac{a_h - a_l}{\sigma} \right) \right]^2 dt \\
&= \frac{\lambda_t}{1 - \pi_t} dt + \frac{\pi_t}{1 - \pi_t} \omega d\tilde{W}_t + \frac{\pi_t^2}{1 - \pi_t} \omega^2 dt = \left( \frac{\lambda_t}{1 - \pi_t} + \omega^2 \pi_t \Phi_t \right) dt + \omega \Phi_t d\tilde{W}_t
\end{align*}
\]

Definition process $\{\tilde{W}_t\}$ as follows:

\[
d\tilde{W}_t = (\omega \pi_t - \sigma) dt + d\tilde{W}_t
\]

And the new measure $P^*$ satisfying:

\[
\frac{dP^*}{dP} = \exp \left\{ -\frac{1}{2} \left( \sigma - \omega \pi_t \right)^2 dt + \int_0^t \left( -\sigma + \omega \pi_t \right) d\tilde{W}_s \right\} = \exp \left\{ \frac{1}{2} \left( \sigma + \omega \pi_t \right)^2 dt + \int_0^t \left( -\sigma + \omega \pi_t \right) d\tilde{W}_s \right\}
\]

where $\omega = \frac{a_h - a_l}{\sigma}$.

According to Girsanov theorem, $\tilde{W}_t$ is a $P^*$-Brownian motion.

We have:

\[
\begin{align*}
d\Phi_t &= \left( \frac{\lambda_t}{1 - \pi_t} + \omega^2 \pi_t \Phi_t \right) dt + \omega \Phi_t d\tilde{W}_t \\
&\Rightarrow \Phi_t = Z_t \left( \Phi_0 + \lambda_t \int_0^t Z_s^{-1} ds \right) \\
&= \exp \left\{ \left( \lambda + \sigma \omega \pi_t \right) t + \sigma \tilde{W}_t \right\}
\end{align*}
\]

\[
\begin{align*}
\frac{dX_t}{X_t} &= \left( 1 - \pi_t \right) a_t + \pi_t a_i \right) dt + \sigma \left( d\tilde{W}_t - (\omega \pi_t - \sigma) dt \right) = \left( a_t + \sigma^2 \right) dt + \sigma d\tilde{W}_t \\
&\Rightarrow X_t = X_0 \exp \left\{ \left( a_t + \sigma^2 \right) t + \sigma \tilde{W}_t \right\}
\end{align*}
\]

Consider the following process:

\[
\eta_t = \exp \left\{ \frac{1}{2} \int_0^t \left( \sigma - \omega \pi_t \right)^2 dt - \int_0^t \left( \sigma - \omega \pi_t \right) d\tilde{W}_s \right\}
\]

\[
= \exp \left\{ \frac{1}{2} \int_0^t \left( \sigma - \omega \pi_t \right)^2 dt - \int_0^t \left( \sigma - \omega \pi_t \right) d\tilde{W}_s \right\} = \exp \left\{ -\frac{1}{2} \int_0^t \left( \sigma - \omega \pi_t \right)^2 dt + \int_0^t \left( \sigma - \omega \pi_t \right) d\tilde{W}_s \right\}
\]

Is a $\mathcal{F}^X$-martingale and \[ \frac{d\eta_t}{\eta_t} = (\omega \pi_t - \sigma) d\tilde{W}_t \] with $\eta_0 = 1$.

Let $U_t$ is process determined by $dU_t = -\lambda_U dt - \sigma U_t d\tilde{W}_t$, we have $U_t = \frac{e^{\left( a_t - \lambda \right)t}}{X_t}$, $U_0 = \frac{1}{X_0}$

Consider the following process:

\[
Y_t = \frac{(1 + \Phi_t) U_t}{(1 - \Phi_0) U_0}, \quad Y_0 = 1.
\]

According to Ito formula we have:
\[\begin{align*}
dY_t &= U_t \left[ (\lambda + (\lambda + \sigma \omega) \Phi_t ) dt + \omega \Phi_t d\tilde{W}_t \right] + (1 + \Phi_t ) \left[ -\lambda U_t dt - \sigma U_t d\tilde{W}_t \right] - \omega \Phi_t \sigma U_t dt \\
&= U_t \omega \Phi_t d\tilde{W}_t - (1 + \Phi_t ) \sigma U_t d\tilde{W}_t = (\Phi_t - (1 + \Phi_t ) \sigma) U_t d\tilde{W}_t = \left( \frac{\omega \Phi_t}{1 + \Phi_t } - \sigma \right) (1 + \Phi_t ) U_t d\tilde{W}_t \\
&= (\omega \pi_t - \sigma) Y_t d\tilde{W} \Rightarrow \frac{dY_t}{Y_t} = (\omega \pi_t - \sigma) d\tilde{W}_t
\end{align*}\]

From this we have \( Y_t = \eta_t \forall t \) (a.s.), then \( e^{\gamma t} \eta_t X_t = e^{\gamma t} Y_t X_t = \frac{1 + \Phi_0}{X_0} e^{(\eta_t - \lambda \gamma - \lambda \gamma^\prime)} (1 + \Phi_t ) \).

Denote \( X_0 = x, \Phi_0 = \phi \) then \( E_{\rho} e^{\gamma t} X_t = E_{\rho} \eta_t e^{\gamma t} X_t = E_{\rho} \eta_t e^{\gamma t} X_t = \frac{x}{1 + \phi} E_{\rho} e^{(\eta_t - \lambda \gamma - \lambda \gamma^\prime)} (1 + \Phi_t ) \).

Problem (2.1) is equivalent to the following problem:

\[ F(\phi) = \inf_{\phi \in \mathbb{R}} E_{\rho} e^{(\eta_t - \lambda \gamma - \lambda \gamma^\prime)} (1 + \Phi_t ) \quad (2.3) \]

Put \( Z_t = e^{(\eta_t - \lambda \gamma - \lambda \gamma^\prime)} (1 + \Phi_t ) \). According to Ito formula we have:

\[ \begin{align*}
dZ_t &= (a_\gamma - \lambda - r) e^{(\eta_t - \lambda \gamma - \lambda \gamma^\prime)} (1 + \Phi_t ) dt + e^{(\eta_t - \lambda \gamma - \lambda \gamma^\prime)} d\Phi_t \\
&= (a_\gamma - \lambda - r) e^{(\eta_t - \lambda \gamma - \lambda \gamma^\prime)} (1 + \Phi_t ) dt + e^{(\eta_t - \lambda \gamma - \lambda \gamma^\prime)} \left[ (\lambda + (\lambda + \sigma \omega) \Phi_t ) dt + \omega \Phi_t d\tilde{W}_t \right] \\
&= e^{(\eta_t - \lambda \gamma - \lambda \gamma^\prime)} \left[ (a_\gamma - \lambda - r) (1 + \Phi_t ) + (\lambda + (\lambda + \sigma \omega) \Phi_t ) \right] dt + e^{(\eta_t - \lambda \gamma - \lambda \gamma^\prime)} \omega \Phi_t d\tilde{W}_t \\
&= e^{(\eta_t - \lambda \gamma - \lambda \gamma^\prime)} \left[ (a_\gamma - r) + (a_\gamma - r) \Phi_t \right] dt + e^{(\eta_t - \lambda \gamma - \lambda \gamma^\prime)} \omega \Phi_t d\tilde{W}_t
\end{align*}\]

So, the drift of \( dU_t \) is \( e^{(\eta_t - \lambda \gamma - \lambda \gamma^\prime)} \left[ (a_\gamma - r) + (a_\gamma - r) \Phi_t \right] \), it will be positive if \( (a_\gamma - r) + (a_\gamma - r) \Phi_t > 0 \Leftrightarrow \Phi_t > \frac{r - a_\gamma}{a_\gamma - r} \), and it will be negative if \( \Phi_t < \frac{r - a_\gamma}{a_\gamma - r} \). This suggests to us that the optimal stopping time is the first time the process hits \( \Phi_t \) in area \([A, \infty)\) with some \( A \). Moreover pair \((A, F)\) satisfy the following free boundary problem:

\[\begin{align*}
\mathcal{L} F + (a_\gamma - \lambda - r) F &= 0 \quad 0 < z < A \\
F(z) &= 1 + z \quad z \geq A \\
F'(A) &= 1 \\
F(0 +) &= \infty
\end{align*}\]

(2.4)

where \( \mathcal{L} \) is infinitesimal operator \( \mathcal{L} F = \frac{w^2 z^2}{2} F'' + (\lambda + (\lambda + \omega \sigma) z) F' \).

Differential Equation in (2.4) has the general solution as follows:

\[ F(z) = C_1 F_1(z) + C_2 F_2(z) \]

where

\[ F_1(z) = x \sqrt{-\omega^2 + 2 \omega \sigma + 2 \lambda} \sqrt{\frac{2 \lambda}{\omega^2}} \text{ Whittaker } W \left( \frac{-\omega^2 + \sigma \omega + \lambda}{\omega^2} , -\frac{2 \lambda}{\omega^2} \right) \]

\[ -4 \omega^2 \lambda - 4 \omega^2 (a_\gamma + a_\gamma - 2 r) + \omega^2 + 4 (a_\gamma - a_\gamma + \lambda)^2 \]

\[ \frac{2 \omega^2}{2 \omega^2} \]

\[ F_2(z) = x \sqrt{-\omega^2 + 2 \omega \sigma + 2 \lambda} \sqrt{\frac{2 \lambda}{\omega^2}} \text{ Whittaker } M \left( \frac{-\omega^2 + \sigma \omega + \lambda}{\omega^2} , -\frac{2 \lambda}{\omega^2} \right) \]

\[ -4 \omega^2 \lambda - 4 \omega^2 (a_\gamma + a_\gamma - 2 r) + \omega^2 + 4 (a_\gamma - a_\gamma + \lambda)^2 \]

\[ \frac{2 \omega^2}{2 \omega^2} \]
and

\[ M(a, b, z) = e^{-\frac{z^2}{2}} \frac{1}{\sqrt{\pi}} \int_0^{\infty} e^{\frac{-t}{2}} (1 - t)^{b-a-1} \, dt \]

Whittaker \( W(a, b, z) = e^{-\frac{z^2}{2}} \frac{1}{\sqrt{\pi}} \int_0^{\infty} e^{\frac{-t}{2}} (1 + t)^{b-a-1} \, dt \)

Changing variables and using some analytic transformations we obtain

\[ F_1(z) = \int_0^{\infty} e^{\alpha u} u^{(\beta + \gamma - 3)/2} (1 + zu)^{(\gamma - \beta + 1)/2} \, du \]

and \( F_2(z) = \int_0^{\infty} e^{\alpha u} u^{(\beta + \gamma - 3)/2} (1 - zu)^{(\gamma - \beta + 1)/2} \, du \)

Denote

\[ \alpha = \frac{2\gamma}{\omega} > 0, \beta = \alpha + \frac{2\sigma}{\omega} > 0, \gamma = \sqrt{\beta - 1} \frac{8(\gamma + r - a)}{\omega^2} > |\beta - 1|. \]

We have

\[ F_1(z) = \int_0^{\infty} e^{\alpha u} u^{(\beta + \gamma - 3)/2} (1 + zu)^{(\gamma - \beta + 1)/2} \, du \]

and \( F_2(z) = \int_0^{\infty} e^{\alpha u} u^{(\beta + \gamma - 3)/2} (1 - zu)^{(\gamma - \beta + 1)/2} \, du \)

Now we consider the property of the function \( F_2(z) \).

First, we calculate the derivative of \( F_2(z) \).

\[ F_2'(z) = \lim_{h \to 0} \frac{F_2(z + h) - F_2(z)}{h} \]

\[ = \lim_{h \to 0} \left( \int_0^{1/h} e^{\alpha u} u^{(\beta + \gamma - 3)/2} (1 - (z + h)u)^{(\gamma - \beta + 1)/2} \, du - \int_0^{1/h} e^{\alpha u} u^{(\beta + \gamma - 3)/2} (1 - zu)^{(\gamma - \beta + 1)/2} \, du \right) \]

\[ = \lim_{h \to 0} \left( \int_0^{1/h} e^{\alpha u} u^{(\beta + \gamma - 3)/2} \left( (1 - (z + h)u)^{(\gamma - \beta + 1)/2} - (1 - zu)^{(\gamma - \beta + 1)/2} \right) \, du \right) \]

\[ = \lim_{h \to 0} \left( \int_0^{1/h} e^{\alpha u} u^{(\beta + \gamma - 3)/2} \left( (1 - (z + h)u)^{(\gamma - \beta + 1)/2} - (1 - zu)^{(\gamma - \beta + 1)/2} \right) \, du \right) \]
Because $\gamma = \sqrt{(\beta - 1)^2 + \frac{8(\lambda + r - a_i)}{\omega^2}} > \beta - 1$ so $F'_2(z) < 0$.

According to average integral theorem we have $F'_2(z) = e^{u^*} \left(1 - zu^*\right)^{(\beta - 1)/2} \int_0^1 u \left(\beta + \gamma - 1\right)^{(\beta - 1)/2} du$

where $u^* \in \left(0, \frac{1}{z}\right)$. Moreover, with small enough $z$ we have $zu^* < \sqrt{z} \Leftrightarrow z < \frac{1}{\left(u^*\right)^2}$. Then we have:

$F'_2(z) > \left(1 - \sqrt{z}\right) \left(\beta - 1\right)^{(\beta - 1)/2} \int_0^1 u \left(\beta + \gamma - 1\right)^{(\beta - 1)/2} du = \frac{2}{\beta + \gamma - 1} \left(1 - \sqrt{z}\right) \left(\beta - 1\right)^{(\beta - 1)/2} \int_0^1 u \left(\beta - 1\right)^{(\beta - 1)/2} du = \frac{\gamma - \beta + 1}{2} \int_0^1 e^{-at}(1 + zt) \left(\beta - 1\right)^{(\beta - 1)/2} dt > 0$ because $r > a_i$ so

$\gamma = \sqrt{(\beta - 1)^2 + \frac{8(\lambda + r - a_i)}{\omega^2}} > \beta - 1$

And then $F'_2(z)$ is increasing function.

Moreover $F'_2(z) = \frac{\gamma - \beta + 1}{2} \int_0^1 e^{-at}(1 + zt) \left(\beta - 1\right)^{(\beta - 1)/2} dt < 0$ because

$\gamma = \sqrt{(\beta - 1)^2 + \frac{8(\lambda + r - a_i)}{\omega^2}} < \sqrt{(\beta - 1)^2 + \frac{8(\lambda + a_i - a_i)}{\omega^2}}$

So $\gamma < \sqrt{(\beta - 1)^2 + \frac{4\lambda + 2 a_i}{\omega^2}} = \sqrt{(\beta - 1)^2 + 4\beta} = \beta + 1$ and $F'_2(z) < 0$ means that the function $F'_2(z)$ is increasing and convex function on $(0, \infty)$. Figure 1 shows the graph of function $F'_2(z)$, we can check the increase and convex properties of it. The graph of $F'_2(z)$ is shown in Figure 2, we can see that it tends to infinite when $z$ as $0+$.

But $F'(0+) < \infty$, then $C_2 = 0$ so $F(z) = C_3 \int_0^1 e^{-at}(1 + zu) \left(\beta - 1\right)^{(\beta - 1)/2} du$

By (2.4) we have

$\left\{\begin{array}{l}
C_3 \int_0^1 e^{-at}(1 + zu) \left(\beta - 1\right)^{(\beta - 1)/2} du = 1 + A \\
C_3 \frac{\gamma - \beta + 1}{2} \int_0^1 e^{-at}(1 + zu) \left(\beta - 1\right)^{(\beta - 1)/2} du = 1
\end{array}\right.$

So $A$ is the solution of equation

$2 \int_0^1 e^{-at}(1 + zu) \left(\beta - 1\right)^{(\beta - 1)/2} du = (1 + A)(\gamma - \beta + 1) \int_0^1 e^{-at}(1 + zu) \left(\beta - 1\right)^{(\beta - 1)/2} dt$ (2.5)
Lemma 2.1. Equation (2.5) has unique solution.

Proof: Equation (2.5) is equivalent to

$$2 \int_0^{\infty} e^{-au} u^{(\beta + \gamma - 3)/2} (1 + xu)^{(\gamma - \beta + 1)/2} \, du - (1 + x)(\gamma - \beta + 1) \int_0^{\infty} e^{-au} u^{(\beta + \gamma - 1)/2} (1 + xu)^{(\gamma - \beta - 1)/2} \, du = 0$$

Put: $$f(x) = 2 \int_0^{\infty} e^{-au} u^{(\beta + \gamma - 3)/2} (1 + xu)^{(\gamma - \beta + 1)/2} \, du - (1 + x)(\gamma - \beta + 1) \int_0^{\infty} e^{-au} u^{(\beta + \gamma - 1)/2} (1 + xu)^{(\gamma - \beta - 1)/2} \, du$$

We have $$f(0+) = 2 \int_0^{\infty} e^{-au} u^{(\beta + \gamma - 3)/2} \, du - (\gamma - \beta + 1) \int_0^{\infty} e^{-au} u^{(\beta + \gamma - 1)/2} \, du$$

For integrals

$$I_1 = 2 \int_0^{\infty} e^{-au} u^{(\beta + \gamma - 3)/2} \, du = \frac{4}{\beta + \gamma - 1} \int_0^{\infty} e^{-au} u^{(\beta + \gamma - 1)/2} \, du = \frac{4}{\beta + \gamma - 1} \left[ e^{-au} u^{(\beta + \gamma - 1)/2} \bigg|_0^{\infty} + \alpha \int_0^{\infty} e^{-au} u^{(\beta + \gamma - 1)/2} \, du \right]$$

$$= \frac{4\alpha}{\beta + \gamma - 1} \int_0^{\infty} e^{-au} u^{(\beta + \gamma - 1)/2} \, du$$
f(0+) = \left[ \frac{4\alpha}{\beta - 1} + (\beta - 1 - \gamma) \right] e^{-\alpha u(\beta + y - 1)/2} du = \left[ \frac{4\alpha + (\beta - 1 - \gamma)^2}{\beta - 1 + \gamma} \right] e^{-\alpha u(\beta + y - 1)/2} du

So:

\begin{align*}
&= \left[ \frac{8\lambda}{\omega^2} - \frac{8(\lambda + r - a)}{\omega^2} \right] e^{-\alpha u(\beta + y - 1)/2} du = \frac{-8(r - a)}{\omega^2 (\beta - 1 + \gamma)} e^{-\alpha u(\beta + y - 1)/2} du < 0
\end{align*}

\begin{align*}
f'(x) &= (\gamma - \beta + 1) e^{-\alpha u(\beta + y - 1)/2} (1 + xu)^{(\beta - 1)/2} du - (\gamma - \beta + 1) e^{-\alpha u(\beta + y - 1)/2} (1 + xu)^{(\beta - 1)/2} du \\
&= -(1 + x) \frac{\gamma - \beta - 1}{2} (\gamma - \beta + 1) e^{-\alpha u(\beta + y - 1)/2} (1 + xu)^{(\beta - 1)/2} du
\end{align*}

Because

\begin{align*}
\gamma < \sqrt{(\beta - 1)^2 + \frac{4 \times 2(\lambda + \sigma \omega)}{\omega^2}} = \sqrt{(\beta - 1)^2 + 4\beta} = \beta + 1; \quad \gamma = \sqrt{(\beta - 1)^2 + \frac{8(\lambda + r - a)}{\omega^2}} > \beta - 1
\end{align*}

and

so \( f'(x) = -(1 + x) \frac{\gamma - \beta - 1}{2} (\gamma - \beta + 1) e^{-\alpha u(\beta + y - 1)/2} (1 + xu)^{(\beta - 1)/2} du > 0, \forall x > 0 \)

We will prove \( \lim_{x \to +\infty} f(x) = +\infty \)

Indeed, for large enough \( x \) we have \( 1 + xu \sim xu \) and

\begin{align*}
f(x) &= 2 e^{-\alpha u(\beta + y - 1)/2} (1 + xu)^{(\beta - 1)/2} du - (1 + x)(\gamma - \beta + 1) e^{-\alpha u(\beta + y - 1)/2} (1 + xu)^{(\beta - 1)/2} du \\
&\sim 2x (\beta - 1)^2 e^{-\alpha u(\beta + y - 1)/2} u^{(\beta - 1)/2} du - (1 + x)(\gamma - \beta + 1) x^{(\beta - 1)/2} e^{-\alpha u(\beta + y - 1)/2} u^{(\beta - 1)/2} du \\
&= 2x (\beta - 1)^2 e^{-\alpha u(\beta + y - 1)/2} u^{(\beta - 1)/2} du - (\gamma - \beta + 1) x^{(\beta - 1)/2} e^{-\alpha u(\beta + y - 1)/2} u^{(\beta - 1)/2} du
\end{align*}

Thus, \( \lim_{x \to +\infty} f(x) < 0, \lim_{x \to +\infty} f(x) = +\infty \) và \( f(x) \) is increasing function so \( f(x) = 0 \) have unique experience. So the theorem is proved.

The graph of \( f(x) \) is shown in Figure 3, it is an increase function and \( \lim_{x \to +\infty} f(x) < 0, \lim_{x \to +\infty} f(x) = +\infty \).

**Theorem 2.2.** Stopping time \( \tau_A = \inf \{ t \geq 0 : \Phi(t) \geq A \} \) is the optimal stopping time for (2.1).

**Proof:**

According to the general theory of optimal stopping time is the point of making the expression:

\[(\mathcal{L} + a_L - \lambda - r)(1 + \phi) = a_L - r + (a_L - r)\phi\]

is positive will be under the continuation area: \( C = \{ \phi : F(\phi) > 1 + \phi \} \).

And stopping time will be optimal if it is the first time that the \( \{ \Phi \}_t \) hits the stopping regions \( D = \{ \phi : F(\phi) = 1 + \phi \} \) or \( \tau_A = \inf \{ t \geq 0 : \Phi(t) \geq A \} \).

In Figure 4, we see a simulation of processes: Asset Price process, process \( \Phi(t) \), threshold probability, posterior probability process and the optimal stopping time with

\[a_L = 0.1; a_H = 0.2; \sigma = 0.4 ; r = 0.05; \lambda = 0.3 \]
Figure 3. Graph of the function $f(x)$.

Figure 4. A simulation of asset price process, the posterior probability process, process $\Phi(t)$, the threshold probability and the optimal stopping time (the optimal time to buy).

We find that $\tau^* = 0.63$ and the asset price (discounted) to buy is $1.24 < 1.5 = X(0)$. In Figure 5, we simulate 4 processes: Asset Price process, process $\Phi(t)$, threshold probability posterior probability process and the optimal stopping time with parameters $a_i = -0.1; a_b = 0.2; \sigma = 0.4; r = 0.05; \lambda = 0.3$.

We find that $\tau^* = 0.58$ and the asset price (discounted) to buy is $1.27 < 1.5 = X(0)$. 

140
3. Conclusion

In this paper, we consider the problem of buying an asset when the asset price is modeled by the geometric Brownian motion which has a change point, where price growth rate is the Markov process with two states that describes the decreasing and increasing of asset prices process on the market. For buying problem we assume that the price will be a shift from decreasing to increasing prices and a buying decision is made when the probability of decreasing state surpassed some certain threshold. We have to simulate the price process with a number of parameters and conduct numerical solution to the experimental buying threshold. In the next study we will consider the problem with assumptions that are closer to reality and more constraints.

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References


